

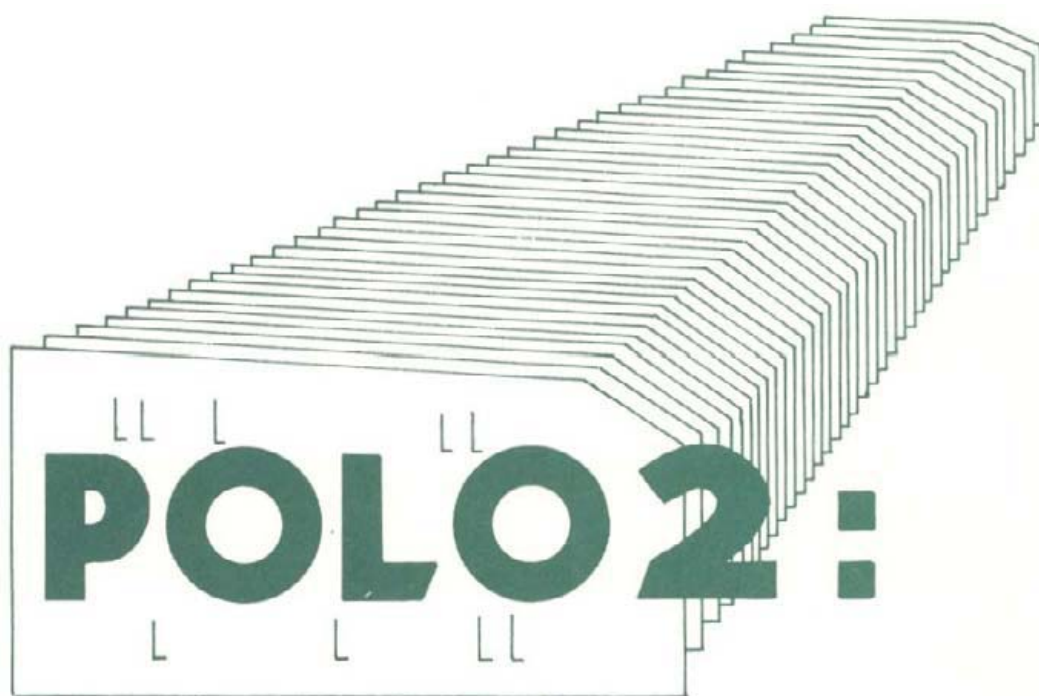


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a user's guide to multiple Probit Or LOGit analysis

Robert M. Russell,

N. E. Savin,

Jacqueline L. Robertson

Authors:

ROBERT M. RUSSELL has been a computer programmer at the Station since 1965. He was graduated from Graceland College in 1953, and holds a B.S. degree (1956) in mathematics from the University of Michigan. **N. E. SAVIN** earned a B.A. degree (1956) in economics and M.A. (1960) and Ph.D. (1969) degrees in economic statistics at the University of California, Berkeley. Since 1976, he has been a fellow and lecturer with the Faculty of Economics and Politics at Trinity College, Cambridge University, England. **JACQUELINE L. ROBERTSON** is a research entomologist assigned to the Station's insecticide evaluation research unit, at Berkeley, California. She earned a B.A. degree (1969) in zoology, and a Ph.D. degree (1973) in entomology at the University of California, Berkeley. She has been a member of the Station's research staff since 1966.

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POLO2:

a user's guide to multiple Probit Or LLogit analysis

Robert M. Russell,

N. E. Savin,

Jacqueline L. Robertson

CONTENTS

Introduction	1
1. General Statistical Features	1
2. Data Input Format	2
2.1 Starter Cards	2
2.2 Title Card	2
2.3 Control Card	3
2.4 Transformation Card	4
2.4.1 Reverse Polish Notation	4
2.4.2 Operators	4
2.4.3 Operands	4
2.4.4 Examples	4
2.5 Parameter Label Card	5
2.6 Starting Values of the Parameters Card	5
2.7 Format Card	5
2.8 Data Cards	5
2.9 End Card	6
3. Limitations	6
4. Data Output Examples	6
4.1 Toxicity of Pyrethrum Spray and Film	6
4.1.1 Models	6
4.1.2 Hypotheses	6
4.1.3 Analyses Required	7
4.1.4 Input	7
4.1.5 Output	9
4.1.6 Hypotheses Testing	19
4.1.7 Comparison with Published Calculations	19
4.2 Vaso-Constriction	19
4.2.1 Models, Hypothesis, and Analyses Required	19
4.2.2 Input	19
4.2.3 Output	20
4.2.4 Hypothesis Testing	21
4.3 Body Weight as a Variable: Higher Order Terms	25
4.3.1 Models and Hypothesis	25
4.3.2 Input	25
4.3.3 Output	26

4.4	Body Weight as a Variable: PROPORTIONAL Option	29
4.4.1	Models and Hypotheses	29
4.4.2	Input	29
4.4.3	Output	30
4.4.4	Hypothesis Testing	30
4.5	Body Weight as a Variable: BASIC Option	30
4.5.1	Input	33
4.5.2	Output	33
5.	Error Messages	36
6.	References	37

Many studies involving quantal response include more than one explanatory variable. The variables in an insecticide bioassay, for example, might be the dose of the chemical as well as the body weight of the test subjects. POLO2 is a computer program developed to analyze binary quantal response models with one to nine explanatory variables. Such models are of interest in insecticide research as well as in other subject areas. For examples of other applications, texts such as those by Domencich and McFadden (1975) and Maddala (1977) should be consulted.

For models in which only one explanatory variable (in addition to the constant) is present, another program, POLO (Russell and others 1977, Savin and others 1977, Robertson and others 1980) is available. However, the statistical inferences drawn from this simple model may be misleading if relevant explanatory variables have been omitted. A more satisfactory approach is to begin the analysis with a general model which includes all the explanatory variables suspected as important in explaining the response of the individual. One may then test whether certain variables can be omitted from the model. The necessary calculations for carrying out these tests are performed by POLO2. If the extra variables are not significant in the multiple regression, a simple regression model may be appropriate.

The statistical documentation of POLO2, descriptions of its statistical features, and examples of its application are described in articles by Robertson and others (1981 a, b), and Savin and others (1981).

The POLO2 program is available upon request to:

Director
Pacific Southwest Forest and Range Experiment Station
P.O. Box 245
Berkeley, California 94701
Attention: Computer Services Librarian

A magnetic tape with format specifications should be sent with the request. The program is currently operational on

the Univac 1100 Series, but can be modified for use with other large scientific computers. The program is not suitable for adaptation to programmable desk calculators.

This guide was prepared to assist users of the POLO2 program. Selected statistical features of the program are described by means of a series of examples chosen from our work and that of others. A comprehensive description of all possible situations or experiments amenable to multivariate analyses is beyond the scope of this guide. For experiments more complex than those described here, a statistician or programmer, or both, should be consulted regarding the appropriate use of POLO2.

1. GENERAL STATISTICAL FEATURES

Consider a sample of I individuals indexed by $i = 1, \dots, I$. For individual i there is an observed $J \times 1$ vector $s_i' = (s_{1i}, \dots, s_{ji})$ of individual characteristics. In a binary quantal response model the individual has two responses or choices. These can be denoted by defining the binomial variable

$f_i = 1$ if the first response occurs
(if alternative 1 is chosen),
 $f_i = 0$ if the second response occurs
(if alternative 2 is chosen).

For example, in a bioassay of toxicants the individuals are insects and the possible responses are dead or alive. The measured characteristics may include the dose of the toxicant, the insect's weight and its age.

The probability (P) that $f_i = 1$ is

$$P_i = F(\beta' z_i)$$

where F is a cumulative distribution function (CDF) mapping points on the real line into the unit interval, $\beta' = (\beta_1, \dots, \beta_K)$ is a $K \times 1$ vector of unknown parameters, $z_{ki} = z_k(s_i)$ is a numerical function of s_i , and $z_i' = (z_{1i}, \dots, z_{Ki})$ is $K \times 1$ vector of these numerical functions. If, for instance, weight is one of the measured characteristics, then the function z_{ki} may be the weight itself, the logarithm of the weight or the square of the weight.

For the probit model

$$P_i = F(\beta' z_i) = \Phi(\beta' z_i)$$

where Φ is the standard normal CDF. For the logit model

$$P_i = F(\beta' z_i) = 1 / [1 + e^{-\beta' z_i}].$$

POLO2 estimates both models by the maximum likelihood (ML) method with grouped as well as ungrouped data.

The ML procedure can be applied to the probability function $P_i = F(\beta' z_i)$ where F is any CDF. Since f_i is a binomial variable, the log of the probability of observing a given sample is

$$L = \sum_{i=1}^I [f_i \log P_i + (1 - f_i) \log(1 - P_i)]$$

where L is referred to as the log likelihood function. The ML method selects as an estimate of β that vector which maximizes L . In other words, the ML estimator for β maximizes the calculated probability of observing the given sample.

When the data are grouped there are repeated observations for each vector of values of the explanatory variables. With grouped data, we change the notation as follows. Now let I denote the number of groups and $i=1, \dots, I$ denote the levels (z_i, s_i) of the explanatory variables. Let n_i denote the number of observations at level i and r_i denote the number of times that the first response occurs. The log likelihood function for grouped data is then

$$L = \sum_{i=1}^I [r_i \log P_i - (n_i - r_i) \log(1 - P_i)]$$

Again, the ML method selects the vector that maximizes the log likelihood function L as an estimate of β . For further discussion of the estimation of probit and logit models with several explanatory variables, see Finney (1971) and Domencich and McFadden (1975).

The maximum of the log likelihood is reported to facilitate hypothesis testing. Two indicators of goodness-

of-fit are routinely calculated. One is the prediction success table (Domencich and McFadden 1975), which compares the results predicted by the multiple regression model with the results actually observed. The other goodness-of-fit indicator is the calculation of the likelihood ratio statistic for testing the hypothesis that all coefficients in the regression are equal to zero. Finally, a general method for transformation of variables is included.

2. DATA INPUT FORMAT

2.1 Starter Cards

Every POLO2 run starts with five cards that call the program from a tape (fig. 1). These cards reflect the current Univac 1100 implementation and would be completely different if POLO2 were modified to run on a different computer. All of the remaining input described in sections 2.2-2.9 would be the same on any computer.

	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	32	33	34	35				
1	@	R	U	N	,	8					I	R	V	,	1	4	3	-	8	8	0	1	,	,	3		1	0	0						P	O	L	O	2
2	@	A	S	G	,	T					X	.	,	T	,	U	2	0	0	0	9																		
3	@	C	O	P	Y	,	G				X	.																											
4	@	F	R	E							X	.																											
5	@	X	Q	T							P	O	L	O	2																								

Figure 1.

Cards 2-5 must be punched as shown. In card 1, the user's identification and account number should be placed in columns 10-21. Column 24 is the time limit in minutes; the page limit is listed in columns 26-28. Both time and page limits may be changed to meet particular needs.

2.2 Title Card

Each data set begins with a title card that has an equal sign (=) punched in column 1. Anything desired may be placed in columns 2-80 (fig. 2). This card is useful in documenting the data, the model, the procedures used for the analysis, or equivalent information. The information is reprinted at the top of every page of the output. Only one title card per data set may be used.

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	32	33	34	35	36	37	38	39	40	41	42	43	44	45	46	47	48	49	50
=WEIGHT STUDY, PYRETHRINS. STANDARD1. LOG(D), LOG(W).																																																	

Figure 2.

2.3 Control Card

The information on this card controls the operation of the program. All items are integers, separated by commas (fig. 3). These numbers need not occur in fixed fields (specific columns on a card). Extra spaces may be inserted before or after the commas as desired (fig. 4). Twelve integers must be present. These are, from left to right:

Position	Internal program designation	Explanation
1	NVAR	Number of regression coefficients in the model, including the constant term, but not including natural response.
2	NV	Number of variables to be read from a data card. This number corresponds to the number of F's and I's on the Format Card (sec. 2.7). Normally, NV=NVAR-1 because the data do not include a constant 1 for the constant term. NV may differ from NVAR-1 when transformations are used in the analysis.
3	LOGIT	LOGIT=0 if the probit model is to be used; LOGIT=1 if the logit model is desired.
4	KONTROL	One of the explanatory variables (for example, dose) will be zero, for the group when a control group is used. KONTROL is the index of that parameter within the sequence of parameters used, with $2 \leq \text{KONTROL} \leq \text{NVAR}$. This limit indicates that any parameter except the constant term may have a control group.
5	ITRAN	ITRAN=1 if variables are transformed and a transformation card will be read. ITRAN=0 if the variables are to be analyzed as is and there will be no transformation card.
6	ISTARV	ISTAR V=1 if starting values of the parameters are to be input; ISTARV=0 if they will be calculated automatically. The ISTARV=1 option should be used only if the automatic method fails.
7	IECHO	IECHO=1 if all data are to be printed back for error checking. If the data have been scrupulously checked, the IECHO=0 option may be used and the data will not be printed back in their entirety. A sample of the data set will be printed instead.
8	NPARN	NPARN=0 if natural response, such as that which occurs without the presence of an insecticide, is present and is to be calculated as a parameter. If natural response is not a parameter, NPARN=1.
9	IGROUP	IGROUP=0 if there is only one test subject per data card; IGROUP=1 if there is more than one (that is, if the data are grouped).
10	NPLL	Number of parallel groups in the data (see sec. 4.1 for an example). NPLL=0 is read as if it were NPLL=1; in other words, a single data set would compose a group parallel to itself.
11	KERNEL	When a restricted model is to be computed, some variables will be omitted from the model and the

regression will be rerun. The KERNEL control instructs the program how many variables to retain. Variables 1 through KERNEL are retained, where $2 \leq \text{KERNEL} \leq \text{NVAR}$; variables KERNEL+1 through NVAR are dropped. Note that variables can be rearranged as desired through the use of transformation or the "T" editing control on the format card (sec. 2.7). When a restricted model is not desired, KERNEL=0.

12 NITER This integer specifies the number of iterations to be done in search of the maximum log likelihood. When NITER=0, the program chooses a suitable value. If starting values are input (ISTARV=1), NITER=0 will be interpreted as no iterations and starting values become the final values. Unless the final values are known and can be used as the starting values, NITER=50 should achieve maximization.

Figure 3

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	
3	,	2	,	0	,	0	,	1	,	0	,	1	,	1	,	0	,	0	,	0	,	0	

Control card specifying three regression coefficients, two variables to be read from a data card, probit model to be used, no controls, a transformation card to be read, and starting values of the parameters to be calculated automatically. Data will be printed back, natural response is not a parameter, there is one subject per data card, no parallel groups, no restricted model will be calculated, and the program will select the number of iterations to be done to find the maximum values of the likelihood function.

Figure 4

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	
3	,	2	,			1	,	2	,			1	,	0	,	0	,	1	,			1	,	0	,	2	,	0

Control card specifying three regression coefficients, two variables to be read from each data card, and the logit model to be used. The second variable defines a control group, a transformation card to be read, starting values of the parameters will be calculated automatically, data will not be printed back, natural response is not a parameter, data are grouped with more than one individual per card, and there are no parallel groups. A restricted model with the first and second variables will be calculated, and the program will select the number of iterations to be done to find the maximum value of the likelihood function. All integers will be read as intended because each is separated from the next by a comma, despite the presence or absence of a blank space.

2.4 Transformation Card

This card contains a series of symbols written in Reverse "Polish" Notation (RPN) (see *sec. 2.4.1*) which defines one or more transformations of the variables. This option is indicated by ITRAN=1 on the control card (*sec. 2.3*). If ITRAN=0 on the control card, the program will not read a transformation card.

2.4.1 Reverse Polish Notation

Reverse Polish Notation is widely used in computer science and in Hewlett-Packard calculators. It is an efficient and concise method for presenting a series of arithmetic calculations without using parentheses. The calculations are listed in a form that can be acted on directed by a computer and can be readily understood by the user.

The central concept of RPN is a "stack" of operands (numbers). The "stack" is likened to a stack of cafeteria trays. We specify that a tray can only be removed from the top of the stack; likewise, a tray can only be put back in the stack at the top. Reverse Polish Notation prescribes all calculations in a stack of operands. An addition operation, for example, calls the top two operands from the stack (reducing the height by two), adds them together, and places the sum back on the stack. Subtraction, multiplication, and division also take two operands from the stack and return one. Simple functions like logarithms and square roots take one operand and return one.

How do numbers get into the stack? Reverse Polish Notation consists of a string of symbols (the operators) and the operands. The string is read from left to right. When an operand is encountered, it is placed on the stack. When an operator is encountered, the necessary operands are removed, and the result is returned to the stack. At the end of the scan, only one number, the final result, remains.

To write an RPN string, any algebraic formula should first be rewritten in linear form.

For example, $\frac{a+b}{c}$ is rewritten $(a+b)/c$.

The operands are written in the order in which they appear in the linear equation. The operators are interspersed in the string in the order in which the stack operates, that is, $ab+c/$. No parentheses are used. When this string is scanned, the following operations occur: (1) a is put on the stack, (2) b is put on the stack, (3) $+$ takes the top two stack items (a, b) and places their sum back on the stack, (4) c is put on the stack, (5) $/$ takes the top stack item (c), divides it into the next item ($a+b$), and places this result back on the stack. In cases where the subtraction operator in an algebraic formula only uses one operand (for example, $-a+b$), a single-operand negative operator such as N can be used. The string is then written $aNb+$.

Once a string has been scanned, a means must exist to begin another stack for another transformation. This is achieved by an operator $=$, which disposes of the final

result. For example $(a+b)/c=d$ becomes the string $ab+c/d=$. The operator $=$ takes two items from the stack and returns none; the result is stored and the stack is empty.

2.4.2 Operators

The operators used in POLO2 are:

Operator	Number of operands	Number of results	Operation
+	2	1	addition
-	2	1	subtraction
*	2	1	multiplication
/	2	1	division
N	1	1	negation
E	1	1	exponentiation (10^x)
L	1	1	logarithm (base 10)
S	1	1	square root
=	2	0	store result

2.4.3 Operands

The operands are taken from an array of values of the variables, that is, $x_1, x_2, x_3, \dots, x_n$. These variables are simply expressed with the subscripts $(1, 2, 3, \dots, n)$; the subscripts are the operands in the RPN string. The symbols in the string are punched one per column, with no intervening blanks. If several new variables are formed by transformations, their RPN strings follow one after another on the card. The first blank terminates the transformations.

The transformations use $x_1, x_2, x_3, \dots, x_n$, to form new variables that must replace them in the same array. To avoid confusion, the x array is copied into another array, y . A transformation then uses operands from x and stores the result in y . Finally, the y is copied back into x and the transformations are done.

The first NPLL numbers in the x array are dummies, or the constant term (1.0) if NPLL=1. Transformations, therefore, are done on $x_2, x_3, \dots, x_n$, and not on the constant term or the dummy variables.

2.4.4 Examples

Several examples of transformations from algebraic notation to RPN are the following:

Algebraic Notation	RPN
$\log(x_2/x_3) = y_2$	23/L2=
$(x_2+x_3)(x_4+x_5) = y_5$	23+45+*5=
$(x_2)^2 + 2x_2x_3 + (x_3)^2 = y_2$	22*23*+23*+33*+2=
$(x_2)^2 = y_2$ $(x_2)^4 = y_3$	22*2= 22*22**3=
$-x_3 + (x_3)^2 - x_2x_4 = y_2$	3N33*24*-S+2=
$x_2(10x^3 + 10x^4) = y_2$	23E4E+*2=

For another example, let the variables in a data set be x_1, x_2, x_3 , and x_4 ; x_1 is the constant term. We require the transformations $y_2 = \log(x_2/x_3)$, $y_3 = \log(x_3)$, $y_5 = (x_2)^2$, and $y_6 = x_2x_3$; x_4 is left unchanged. The RPN strings are 23/L2=, 3L3=, 22*5=, and 23*6=. The appropriate transformation card for this series is shown in *figure 5*.

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20
2	3	/	L	2	=	3	L	3	=	2	2	*	5	=	2	3	*	6	=

Figure 5.

2.5 Parameter Label Card

This card contains the descriptive labels for all NVAR parameters. These labels are used on the printout. The parameters include the constant term or dummy variables, other explanatory variables, and natural response, if it is present. Labels apply to the explanatory variables after any transformations. Each label is 8 characters long. Use columns 1-8 for the first label, columns 9-16 for the second, and so on (*fig. 6*). If a label does not fill the 8 spaces, begin the label in the leftmost space available (columns 1, 9, 17, 15, 33, and so on).

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31
CONSTANT DOSE														WEIGHT										NATURAL						

Figure 6.

2.6 Starting Values of the Parameters Card

This card is used only under special circumstances, such as quickly confirming calculations in previously published experiments. If this card is to be used, ISTARV=1 on the control card (*sec. 2.3*). The parameters are punched in 10F8.6 format (up to 10 fields of 8 columns each); in each field there is a number with six digits to the right of the decimal point and two to the left. The decimal point need not be punched. The parameters on this card are the same in number and order as on the label card (*fig. 6, 7*).

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22
	3	4	6	7	4				6	6	2	9	2				5	8	8	4	2

Figure 7.

In this example, the constant is 3.4674, β_1 is 6.6292, and β_2 is 5.8842. All POLO2 calculations will be done on the basis of these parameter values if ISTARV=1 and NITER=0 on the control card (*sec. 2.3*).

2.7 Format Card

This card contains a standard FORTRAN format statement with parentheses but without "FORMAT" punched on the card (*fig. 8*). This statement instructs the program how to read the data from each card. A variable occupies specific columns—a field—on each data card; this

area is specified by "F" followed by a number giving the field width. After that number is a decimal point and another number telling where the decimal point is located, if it is not punched on the data cards. For example, "F7.2" means that the data item requires a 7-column field; a decimal point occurs between columns 5 and 6. "F7.0" means whole numbers. When a decimal point is actually punched on a data card, the computer ignores what the format card might say about its location. (For more information, see any FORTRAN textbook.)

Besides the variables, the other items on a data card, such as the group number (K), number of subjects (N), and number of subjects responding (M) must be specified on the format card in "I" (integer) format. Formal editing controls "X" and "T" may be used to skip extraneous columns on a data card or to go to a particular column. For example, "3X" skips 3 columns and "16T" resets the format scan to column 16 regardless of where the scan was previously. All steps in the format statement are separated by commas, and the statement is enclosed in parentheses.

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22
(	1	0	X	,	F	4	.	2	,	F	5	.	1	,	T	2	4	,	I	1	)

Figure 8.

Format card instructing program to skip the first 10 columns of each data card, read the first variable within the next 4 columns assuming 2 decimal places, read the second variable within the next 5 columns assuming 1 decimal place, then go to column 24 and read a single integer, M (M=1 for response; M=0 for no response).

2.8 Data Cards

Punch one card per subject or per group of subjects grouped at identical values of one of the independent variables. All individuals treated with the same dose of an insecticide, for example, might be grouped on a single card, or each might have its own data card. Values of the NV variables are punched, followed by N (the number of subjects) and M (the number responding). If there is only one subject per card (IGROUP=0) (*see sec. 2.3*), N should be omitted.

If parallel groups are being compared (NPLL > 1), the data card must also contain the group number K (K = 1,2,3,...,NPLL) punched before the variables. In summary, a data card contains K, $x_1, x_2, x_3, \dots, x_{NV}, N, M$ with K omitted when NPLL = 0 or 1, and N omitted if IGROUP=0.

Figures illustrating these alternatives will be provided in the examples to follow. If data have already been punched, but are to be used in a different order, the order may be altered in the format card by use of the "T" format editing control. This control will permit the scan to jump backwards, as necessary.

2.9 END Card

To indicate the end of a problem if more problems are to follow in the same job, "END" should be punched in columns 1-3 (fig. 9). If only one problem is analyzed, this card is not necessary.

3. LIMITATIONS

No more than 3000 test subjects may be included in a single analysis. This counts all subjects in grouped data. Including the constant term(s), no more than nine explanatory variables may be used.

4. DATA OUTPUT EXAMPLES

Examples illustrating POLO2 data output and uses of the program's special features for hypotheses testing follow. Each problem is presented in its entirety, from data input through hypotheses testing, with statistics from the output.

4.1 Toxicity of Pyrethrum Spray and Film

Data from experiments of Tattersfield and Potter (1943) are used by Finney (1971, p. 162-169) to illustrate the calculations for fitting parallel probit planes. Insects (*Tribolium castaneum*) were exposed to pyrethrum, a botanical insecticide, either as a direct spray or as a film deposited on a glass disc. We use these data to illustrate multivariate analysis of grouped data, use of dummy variables, use of transformations, the likelihood ratio test for parallelism of probit planes, and the likelihood ratio test for equality of the planes.

4.1.1 Models

The probit model expressing the lethal effect of pyrethrum spray is

$$y_s = \alpha_s + \beta_{s1}x_1 + \beta_{2s}x_2 \quad [1]$$

where y_s is the probit of percent mortality, x_1 is the concentration of pyrethrum in mg/ ml, and x_2 is the weight (deposit) in mg/ cm². The regression coefficients are α_s for the constant, β_{1s} for spray concentration, and β_{2s} for weight. Similarly, the model for the lethal effect of pyrethrum film is

$$y_f = \alpha_f + \beta_{1f}x_1 + \beta_{2f}x_2 \quad [2]$$

where y_f is the probit of percent mortality, x_1 is concentration, and x_2 is weight. The regression coefficients are α_f for the constant, β_{1f} for concentration, and β_{2f} for weight (deposit) of pyrethrum in the film.

4.1.2 Hypotheses

The likelihood ratio (LR) procedure will be used to test three hypotheses. These hypotheses are that the spray and film planes are parallel, that the planes are equal given the assumption that the planes are parallel, and that the planes are equal with no assumption of parallelism.

The LR test compares two values of the logarithm of the likelihood function. The first is the maximum value of the log likelihood when it is maximized unrestrictedly. The second is the maximum value when it is maximized subject to the restrictions imposed by the hypothesis being tested. The unrestricted maximum of the log likelihood is denoted by $L(\Omega)$ and the restricted maximum by $L(\omega)$.

The hypothesis of parallelism is $H(P): \beta_{1s} = \beta_{1f}, \beta_{2s} = \beta_{2f}$. Let L_s and L_f denote the maximum value of the log likelihood for models [1] and [2], respectively. The value $L(\Omega)$ is the sum of L_s and L_f .

- (i) L_s : = ML estimation of [1].
 L_f : = ML estimation of [2].
 $L(\Omega) = L_s + L_f$.

The model with the restrictions imposed is

$$y = \alpha_s x_s + \alpha_f x_f + \beta x_1 + \beta_2 x_2 \quad [3]$$

In this restricted model, dummy variables are used. The dummy variables x_s and x_f are defined as follows:

$$\begin{aligned} x_s &= 1 \text{ for spray; } x_f = 1 \text{ for film;} \\ x_s &= 0 \text{ for film; } x_f = 0 \text{ for spray;} \end{aligned}$$

The value $L(\omega)$ is obtained by estimating [3] by ML.

- (ii) $L(\omega)$: ML estimation of [3].

When H is true, asymptotically,

$$LR = 2[L(\Omega) - L(\omega)] \sim \chi^2(2).$$

In other words, for large samples the LR test statistic has approximately a chi-square distribution with 2 degrees of freedom (df). The df is the number of restrictions imposed by the hypothesis, which in this situation equals the number of parameters constrained to be the same. The LR test accepts $H(P)$ at significance level α if

$$LR \leq \chi_{\alpha}^2(n)$$

where $\chi^2(n)$ denotes the upper significance point of a chi-square distribution with n df.

The hypothesis of equality given parallelism is $H(E|P): \alpha_s = \alpha_f$. Now the unrestricted model is [3] and the restricted model is

$$y = \alpha + \beta_1 x_1 + \beta_2 x_2 \quad [4]$$

Figure 9—Continued

skip the next column, read two 5-digit fields each of which includes a decimal point, then read two 4-column fields of integers. The data cards (*fig. 9-C-5*) are followed by the natural response data card (*fig. 9-C-5a*), then the END card (*fig. 9-C-5b*).

4.1.5 Output

The output for the five analyses is shown in *figure 10*. Except where noted, each analysis is shown in its entirety. The title of the analysis is printed as the first line on each page (*fig. 10-1*, lines 1 and 56; *fig. 10-2*, lines 94 and 149; *fig. 10-3*, lines 184, 239, 297; *fig. 10-4*, lines 324, 378, and 435; *fig. 10-5*, lines 459, 514, and 573). Next, each control card is listed (*fig. 10-1*, line 2; *fig. 10-2*, line 95; *fig. 10-3*, line 185; *fig. 10-4*, line 325; *fig. 10-5*, line 460). The subsequent section of the printout describes the specifications of the analysis and reflects the information on the control card (*fig. 10-1*, lines 3-11; *fig. 10-2*, lines 96-104; *fig. 10-3*, lines 186-195; *fig. 10-4*, lines 326-334; *fig. 10-5*, lines 461-471).

Transformations are reproduced in RPN, just as they were punched on the transformation card (*fig. 10-1*, line 12; *fig. 10-2*, line 105; *fig. 10-3*, line 196; *fig. 10-4*, line 335; *fig. 10-5*, line 472). Parameter labels are reproduced next (*fig. 10-1*, lines 13-16; *fig. 10-2*, lines 106-109; *fig. 10-3*, lines 197-201; *fig. 10-4*, lines 336-339; *fig. 10-5*, lines 473-478), followed by the format statement (*fig. 10-1*, line 17; *fig. 10-2*, line 110; *fig. 10-3*, line 202; *fig. 10-4*, line 340; *fig. 10-5*, line 479).

In the next section, input data are listed as punched on the data cards (*fig. 10-1*, lines 18-30; *fig. 10-2*, lines 111-123; *fig. 10-3*, lines 203-227; *fig. 10-4*, lines 341-363; *fig. 10-5*, lines 480-505). The transformed data are listed after the data input. For the purposes of computation of the prediction success table and other statistics, grouped data are now listed as individual cases. This section of the output has been abbreviated in the figure (*fig. 10-1*, lines 31-65; *fig. 10-2*, lines 124-155; *fig. 10-3*, lines 228-292; *fig. 10-4*, lines 364-429; *fig. 10-5*, lines 506-570). At the end of the transformed data, the total number of cases (observations plus controls) is summarized (*fig. 10-1*, line 66; *fig. 10-2*, line 156; *fig. 10-3*, line 293; *fig. 10-4*, line 430; *fig. 10-5*, line 571). Proportional control mortality is also printed (*fig. 10-5*, line 572).

Initial estimates of the parameters that were computed by the program are printed (*fig. 10-1*, lines 67-68; *fig. 10-2*, lines 157-158; *fig. 10-3*, lines 294-295; *fig. 10-4*, lines 431-432; *fig. 10-5*, lines 574-577), followed by the initial value of the log likelihood (*fig. 10-1*, line 69; *fig. 10-2*, line 159; *fig. 10-3*, line 296; *fig. 10-4*, line 433; *fig. 10-5*, line 578). The program begins with logit iterations that are computationally simpler, then switches to probit calculations. The number of iterations of each type are listed (*fig. 10-1*, lines 70-71; *fig. 10-2*, lines 160-161; *fig. 10-3*, lines 298-299; *fig. 10-4*, lines 434, 436; *fig. 10-5*, lines 579-580) preceding the final value of the log likelihood (*fig. 10-1*, line 72; *fig. 10-2*, line 162; *fig. 10-3*, line 300; *fig. 10-4*, line 437; *fig. 10-5*, line 581).

Parameter values, their standard errors, and their t-ratios (parameter values divided by its standard error) are then presented (*fig. 10-1*, lines 73-76; *fig. 10-2*, lines 163-166; *fig. 10-3*, lines 301-305; *fig. 10-4*, lines 438-441; *fig. 10-5*, lines 582-587). The t-ratios are used to test the significance of each parameter in the regression. The hypothesis that a regression coefficient is zero is rejected at the $\alpha = 0.05$ significance level when the absolute values of the t-ratio is greater than $t = 1.96$, that is, the upper $\alpha = 0.05$ significance point of a t distribution with ∞ df. All parameters in each of the five analyses were significant in this example. These statistics are followed by the covariance matrix for the analysis (*fig. 10-1*, lines 77-81; *fig. 10-2*, lines 167-171; *fig. 10-3*, lines 306-311; *fig. 10-4*, lines 442-446; *fig. 10-5*, lines 588-594).

A prediction success table (*fig. 10-1*, lines 82-91; *fig. 10-2*, lines 172-181; *fig. 10-3*, lines 312-321; *fig. 10-4*, lines 447-456; *fig. 10-5*, lines 595-604) lists the number of individual test subjects that were predicted to be alive and were actually alive, predicted to be alive but were actually dead, predicted to be dead and were actually dead, and predicted to be dead but were actually alive. The numbers are calculated by using maximum probability as a criterion (Domencich and McFadden 1975). The percent correct prediction, rounded to the nearest percent, is calculated as the number predicted correctly (that is, alive when predicted alive, or dead when predicted dead) divided by the total number predicted in that category, times 100. In the pyrethrum spray experiment (*fig. 10-1*), for example, 89 individuals were correctly forecast as alive, but 26 others were dead when they had been predicted to be alive. The correct percent alive is

$$\frac{89}{89 + 26} \times 100,$$

which is 77 percent (*fig. 10-1*, line 90). The overall percent correct (OPC) equals the total number correctly predicted divided by the total number of observations, times 100, rounded to the nearest percent. In the pyrethrum spray example, OPC equals

$$\frac{89 + 195}{333} \times 100,$$

or 85 percent (*fig. 10-1*, line 91). On the basis of random choice, the correct choice should be selected for about 50 percent of the observations. In the five analyses, OPC values indicate reliable prediction by the models used, since all were greater than 80 percent.

The last portion of the output tests the significance of the regression coefficients in each regression (*fig. 10-1*, lines 92-93; *fig. 10-2*, lines 182-183; *fig. 10-3*, lines 322-323; *fig. 10-4*, lines 457-458; *fig. 10-5*, lines 605-606). The hypothesis tested is that all the regression coefficients equal zero. This hypothesis implies that the probability of death is 0.5 at all spray concentrations and film deposits. The log likelihood

$L(\omega)$ is calculated for this restricted model and compared to the maximized log likelihood $L(\Omega)$ for the unrestricted model. When the hypothesis is true, asymptotically,

$$LR = 2[L(\Omega) - L(\omega)] \sim \chi^2(df)$$

where df equals the number of parameters in the

unrestricted model. The hypothesis is accepted at the α level of significance if

$$LR \leq \chi_{\alpha}^2(df).$$

In each of the five, analyses in this example, the hypothesis was rejected at the $\alpha = 0.05$ significance level. All regressions were highly significant.

Figure 10-1

1	PYR S NONAT	POL02	26 AUG 80 14:06:40	4	1
2	CONTROL CARD: 3,2,0,0,1,0,1,1,1,0,0				
3	NUMBER OF VARIABLES: 3				
4	NUMBER OF VARIABLES ON INPUT CARDS: 2				
5	NUMBER OF PARAMETERS: 3				
6	TYPE OF MODEL: PROBIT				
7	THERE ARE TRANSFORMATIONS				
8	INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY				
9	ALL RAW DATA WILL BE PRINTED				
10	NATURAL RESPONSE IS NOT A PARAMETER				
11	MULTIPLE SUBJECTS PER DATA CARD				
12	TRANSFORMATIONS: 2L2=3L3=				
13	PARAMETER LABELS				
14	1 CONSTANT				
15	2 LOG(C)				
16	3 LOG(W)				
17	FORMAT: (1X,1X,2F5.0,2I4)				
18	INPUT DATA				
19	1 5.	2.9	27	1	
20	1 10.	2.9	29	15	
21	1 20.	2.9	30	27	
22	1 40.	2.9	28	28	
23	1 5.	5.7	29	4	
24	1 10.	5.7	29	19	
25	1 20.	5.7	27	26	
26	1 40.	5.7	30	30	
27	1 5.	10.8	30	6	
28	1 10.	10.8	24	15	
29	1 20.	10.8	31	31	
30	1 40.	10.8	19	19	
31	TRANSFORMED DATA				
32	1.	1	1.00000	.698970	.462398
33	2.	0	1.00000	.698970	.462398
34	3.	0	1.00000	.698970	.462398
35	4.	0	1.00000	.698970	.462398
36	5.	0	1.00000	.698970	.462398
37	6.	0	1.00000	.698970	.462398
38	7.	0	1.00000	.698970	.462398
39	8.	0	1.00000	.698970	.462398
40	9.	0	1.00000	.698970	.462398
41	10.	0	1.00000	.698970	.462398
42	11.	0	1.00000	.698970	.462398
43	12.	0	1.00000	.698970	.462398
44	13.	0	1.00000	.698970	.462398
45	14.	0	1.00000	.698970	.462398
46	15.	0	1.00000	.698970	.462398
47	16.	0	1.00000	.698970	.462398
48	17.	0	1.00000	.698970	.462398
49	18.	0	1.00000	.698970	.462398
50	19.	0	1.00000	.698970	.462398
51	20.	0	1.00000	.698970	.462398
52	21.	0	1.00000	.698970	.462398
53	22.	0	1.00000	.698970	.462398
54	23.	0	1.00000	.698970	.462398
55	24.	0	1.00000	.698970	.462398
56	PYR S NONAT	POL02	26 AUG 80 14:06:43	5	7
57	325.	1	1.00000	1.60206	1.03342
58	326.	1	1.00000	1.60206	1.03342
59	327.	1	1.00000	1.60206	1.03342
60	328.	1	1.00000	1.60206	1.03342
61	329.	1	1.00000	1.60206	1.03342
62	330.	1	1.00000	1.60206	1.03342
63	331.	1	1.00000	1.60206	1.03342

Figure 10-1—Continued

```

64 332. 1 1.00000 1.60206 1.03342
65 333. 1 1.00000 1.60206 1.03342
66 333 OBSERVATIONS + 0 CONTROLS = 333 TOTAL
67 INITIAL ESTIMATES OF THE PARAMETERS
68 -4.327488 CONSTANT 3.934193 LOG(C) .7871704 LOG(W)
69 STARTING LIKELIHOOD = -103.9408
70 7 LOGIT EVALUATIONS
71 5 PROBIT EVALUATIONS
72 MAXIMUM LOG LIKELIHOOD = -101.3851
73
74 PARAMETER STD ERROR T RATIO
75 CONSTANT -5.46446 .644240 -8.48203
76 LOG(C) 4.92555 .484850 10.1589
77 LOG(W) 1.09323 .445611 2.45334
78
79 COVARIANCE MATRIX
80 CONSTANT LOG(C) LOG(W)
81 CONSTANT .4150454 -.2669031 -.2008408
82 LOG(C) -.2669031 .2350799 .5576819-01
83 LOG(W) -.2008408 .5576819-01 .1985690
84
85 PREDICTION SUCCESS TABLE
86
87 PREDICTED OUTCOME
88 -----
89 ACTUAL OUTCOME ALIVE DEAD
90 -----
91 ALIVE 49 25
92 DEAD 26 195
93 -----
94 PERCENT CORRECT 77 89
95 OVERALL PERCENT CORRECT: 85
96
97 TESTING HYPOTHESIS THAT THE 3-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
98 LIKELIHOOD RATIO TEST: CHI-SQUARED = 258.8652 DF=3 TAIL PROBABILITY = .000

```

Figure 10-2

```

94 PYR F NUNAT POL02 26 AUG 80 14:06:47 8 8
95 CONTROL CARD: 3,2,0,0,1,0,1,1,1,0,0
96 NUMBER OF VARIABLES: 3
97 NUMBER OF VARIABLES ON INPUT CARDS: 2
98 NUMBER OF PARAMETERS: 3
99 TYPE OF MODEL: PROBIT
100 THERE ARE TRANSFORMATIONS
101 INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY
102 ALL RAW DATA WILL BE PRINTED
103 NATURAL RESPONSE IS NOT A PARAMETER
104 MULTIPLE SUBJECTS PER DATA CARD
105 TRANSFORMATIONS: 2L2=3L3=
106 PARAMETER LABELS
107 1 CONSTANT
108 2 LOG(C)
109 3 LOG(W)
110 FORMAT: (1X,1X,2F5.0,2I4)
111 INPUT DATA
112 2 5. 2.9 29 3
113 2 10. 2.9 30 10
114 2 20. 2.9 29 24
115 2 40. 2.9 29 29
116 2 5. 5.7 27 4
117 2 10. 5.7 28 14
118 2 20. 5.7 28 27
119 2 40. 5.7 29 29
120 2 5. 10.8 28 8
121 2 10. 10.8 28 17
122 2 20. 10.8 28 26
123 2 40. 10.8 17 17

```

Figure 10-2—Continued

```

124 TRANSFORMED DATA
125 1. 1 1.00000 .698970 .462398
126 2. 1 1.00000 .698970 .462398
127 3. 1 1.00000 .698970 .462398
128 4. 0 1.00000 .698970 .462398
129 5. 0 1.00000 .698970 .462398
130 6. 0 1.00000 .698970 .462398
131 7. 0 1.00000 .698970 .462398
132 8. 0 1.00000 .698970 .462398
133 9. 0 1.00000 .698970 .462398
134 10. 0 1.00000 .698970 .462398
135 11. 0 1.00000 .698970 .462398
136 12. 0 1.00000 .698970 .462398
137 13. 0 1.00000 .698970 .462398
138 14. 0 1.00000 .698970 .462398
139 15. 0 1.00000 .698970 .462398
140 16. 0 1.00000 .698970 .462398
141 17. 0 1.00000 .698970 .462398
142 18. 0 1.00000 .698970 .462398
143 19. 0 1.00000 .698970 .462398
144 20. 0 1.00000 .698970 .462398
145 21. 0 1.00000 .698970 .462398
146 22. 0 1.00000 .698970 .462398
147 23. 0 1.00000 .698970 .462398
148 24. 0 1.00000 .698970 .462398

```

```

149 PYR F NONAT POL02 26 AUG 80 14:06:49 9 14

```

```

150 325. 1 1.00000 1.60206 1.03342
151 326. 1 1.00000 1.60206 1.03342
152 327. 1 1.00000 1.60206 1.03342
153 328. 1 1.00000 1.60206 1.03342
154 329. 1 1.00000 1.60206 1.03342
155 330. 1 1.00000 1.60206 1.03342

```

```

156 330 OBSERVATIONS + 0 CONTROLS = 330 TOTAL

```

```

157 INITIAL ESTIMATES OF THE PARAMETERS
158 -4.009471 CONSTANT

```

```
3.357726 LOG(C)
```

```
.9834336 LOG(W)
```

```

159 STARTING LIKELIHOOD = -124.0496

```

```

160 7 LOGIT EVALUATIONS

```

```

161 6 PROBIT EVALUATIONS

```

```

162 MAXIMUM LOG LIKELIHOOD = -122.4908

```

163	PARAMETER	STD ERROR	1 RATIO
164	CONSTANT -4.77369	.545735	-8.91055
165	LOG(C) 3.99017	.381101	10.4701
166	LOG(W) 1.18636	.395729	2.99792

```

167 COVARIANCE MATRIX

```

168	CONSTANT	LOG(C)	LOG(W)
169	CONSTANT .2870119	-.1689206	-.1474890
170	LOG(C) -.1689206	.1452383	.3136549-01
171	LOG(W) -.1474890	.3136549-01	.1566018

```

172 PREDICTION SUCCESS TABLE

```

173	ACTUAL OUTCOME	PREDICTED OUTCOME	
174		ALIVE	DEAD
175	ALIVE	89	33
176	DEAD	25	183
177	PERCENT CORRECT	78	85
178	OVERALL PERCENT CORRECT: 82		

```

182 TESTING HYPOTHESES THAT THE 4-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
183 LIKELIHOOD RATIO TEST: CHI-SQUARED = 212.4950 DF=3 TAIL PROBABILITY = .000

```


Figure 10-3

```

184      PYR SF NONAT PARALLEL                      POL02          26 AUG 80  14:06:52  11  15

185      CONTROL CARD: 4,2,0,0,1,0,1,1,1,2,0,0

186      NUMBER OF VARIABLES: 6
187      NUMBER OF VARIABLES ON INPUT CARDS: 2
188      NUMBER OF PARAMETERS: 4
189      TYPE OF MODEL: PROBIT
190      THERE ARE TRANSFORMATIONS
191      INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY
192      ALL RAW DATA WILL BE PRINTED
193      NATURAL RESPONSE IS NOT A PARAMETER
194      MULTIPLE SUBJECTS PER DATA CARD
195      NUMBER OF PARALLEL GROUPS: 2

196      TRANSFORMATIONS: 3L3=4L4=

197      PARAMETER LABELS
198      1  SPRAY
199      2  FILM
200      3  LOG(C)
201      4  LOG(W)

202      FORMAT: (I1,I1,2F5.0,2I4)

203      INPUT DATA
204      1 5. 2.9 27 1
205      1 10. 2.9 29 15
206      1 20. 2.9 30 27
207      1 40. 2.9 28 28
208      1 5. 5.7 29 4
209      1 10. 5.7 29 19
210      1 20. 5.7 27 26
211      1 40. 5.7 30 30
212      1 5. 10.8 30 6
213      1 10. 10.8 28 15
214      1 20. 10.8 31 11
215      1 40. 10.8 19 19
216      2 5. 2.9 29 5
217      2 10. 2.9 30 10
218      2 20. 2.9 29 24
219      2 40. 2.9 29 24
220      2 5. 5.7 27 4
221      2 10. 5.7 28 14
222      2 20. 5.7 28 27
223      2 40. 5.7 29 24
224      2 5. 10.8 28 6
225      2 10. 10.8 28 17
226      2 20. 10.8 28 26
227      2 40. 10.8 17 17

228      TRANSFORMED DATA
229      1. 1 1.00000 .000000 .698970 .462398
230      2. 0 1.00000 .000000 .698970 .462398
231      3. 0 1.00000 .000000 .698970 .462398
232      4. 0 1.00000 .000000 .698970 .462398
233      5. 0 1.00000 .000000 .698970 .462398
234      6. 0 1.00000 .000000 .698970 .462398
235      7. 0 1.00000 .000000 .698970 .462398
236      8. 0 1.00000 .000000 .698970 .462398
237      9. 0 1.00000 .000000 .698970 .462398
238      10. 0 1.00000 .000000 .698970 .462398

-----
239      PYR SF NONAT PARALLEL                      POL02          26 AUG 80  14:06:55  13  26

240      611. 0 .000000 1.00000 1.00000 1.03342
241      612. 0 .000000 1.00000 1.00000 1.03342
242      613. 0 .000000 1.00000 1.00000 1.03342
243      614. 0 .000000 1.00000 1.00000 1.03342
244      615. 0 .000000 1.00000 1.00000 1.03342
245      616. 0 .000000 1.00000 1.00000 1.03342
246      617. 0 .000000 1.00000 1.00000 1.03342
247      618. 0 .000000 1.00000 1.00000 1.03342
248      619. 1 .000000 1.00000 1.30103 1.03342
249      620. 1 .000000 1.00000 1.30103 1.03342
250      621. 1 .000000 1.00000 1.30103 1.03342
251      622. 1 .000000 1.00000 1.30103 1.03342
252      623. 1 .000000 1.00000 1.30103 1.03342
253      624. 1 .000000 1.00000 1.30103 1.03342
254      625. 1 .000000 1.00000 1.30103 1.03342
255      626. 1 .000000 1.00000 1.30103 1.03342
256      627. 1 .000000 1.00000 1.30103 1.03342
257      628. 1 .000000 1.00000 1.30103 1.03342
258      629. 1 .000000 1.00000 1.30103 1.03342
259      630. 1 .000000 1.00000 1.30103 1.03342
260      631. 1 .000000 1.00000 1.30103 1.03342
261      632. 1 .000000 1.00000 1.30103 1.03342
262      633. 1 .000000 1.00000 1.30103 1.03342

```

Figure 10-3—Continued

```

263 634. 1 .000000 1.00000 1.30103 1.03342
264 635. 1 .000000 1.00000 1.30103 1.03342
265 636. 1 .000000 1.00000 1.30103 1.03342
266 637. 1 .000000 1.00000 1.30103 1.03342
267 638. 1 .000000 1.00000 1.30103 1.03342
268 639. 1 .000000 1.00000 1.30103 1.03342
269 640. 1 .000000 1.00000 1.30103 1.03342
270 641. 1 .000000 1.00000 1.30103 1.03342
271 642. 1 .000000 1.00000 1.30103 1.03342
272 643. 1 .000000 1.00000 1.30103 1.03342
273 644. 1 .000000 1.00000 1.30103 1.03342
274 645. 0 .000000 1.00000 1.30103 1.03342
275 646. 0 .000000 1.00000 1.30103 1.03342
276 647. 1 .000000 1.00000 1.60206 1.03342
277 648. 1 .000000 1.00000 1.60206 1.03342
278 649. 1 .000000 1.00000 1.60206 1.03342
279 650. 1 .000000 1.00000 1.60206 1.03342
280 651. 1 .000000 1.00000 1.60206 1.03342
281 652. 1 .000000 1.00000 1.60206 1.03342
282 653. 1 .000000 1.00000 1.60206 1.03342
283 654. 1 .000000 1.00000 1.60206 1.03342
284 655. 1 .000000 1.00000 1.60206 1.03342
285 656. 1 .000000 1.00000 1.60206 1.03342
286 657. 1 .000000 1.00000 1.60206 1.03342
287 658. 1 .000000 1.00000 1.60206 1.03342
288 659. 1 .000000 1.00000 1.60206 1.03342
289 660. 1 .000000 1.00000 1.60206 1.03342
290 661. 1 .000000 1.00000 1.60206 1.03342
291 662. 1 .000000 1.00000 1.60206 1.03342
292 663. 1 .000000 1.00000 1.60206 1.03342

293 663 OBSERVATIONS + 0 CONTROLS = 663 TOTAL

294 INITIAL ESTIMATES OF THE PARAMETERS
295 -4.101413 SPRAY -4.204135 FILM 3.631875 LOG(C) .8847998 LOG(W)

296 STARTING LIKELIHOOD = -228.9565

-----
297 PYR SF NONAT PARALLEL PDL02 26 AUG 80 14:07:00 15 27

298 7 LOGIT EVALUATIONS
299 5 PROBIT EVALUATIONS

300 MAXIMUM LOG LIKELIHOOD = -225.1918

301 PARAMETER
302 SPRAY -4.97434 .411879 -12.0772
303 FILM -5.12994 .417825 -12.2777
304 LOG(C) 4.34963 .299519 14.6556
305 LOG(W) 1.12907 .294399 3.83517

306 COVARIANCE MATRIX
307 SPRAY FILM LOG(C) LOG(W)
308 SPRAY .1696447 .1630720 -.1016954 -.8344456-01
309 FILM .1630720 .1745781 -.1039320 -.8417556-01
310 LOG(C) -.1016954 -.1039320 .8971165-01 .1988976-01
311 LOG(W) -.8344456-01 -.8417556-01 .1988976-01 .8667098-01

312 PREDICTION SUCCESS TABLE
313 PREDICTED OUTCOME
314 -----
315 ACTUAL OUTCOME ALIVE DEAD
316 -----
317 ALIVE 178 56
318 DEAD 51 378
319 -----
320 PERCENT CORRECT 78 87
321 OVERALL PERCENT CORRECT: 84

322 TESTING HYPOTHESIS THAT THE 4-ETA-PARAMETER MODEL AND THE 0-ETA-PARAMETER MODEL ARE THE SAME
323 LIKELIHOOD RATIO TEST: CHI-SQUARED = 468.7280 DF=4 TAIL PROBABILITY = .000

```

Figure 10-4

324	PYR SF NONAT EQUALITY	POL02	26 AUG 80	14:07:02	17	28
325	CONTROL CARD: 3,2,0,0,1,0,1,1,1,0,0					
326	NUMBER OF VARIABLES: 3					
327	NUMBER OF VARIABLES ON INPUT CARDS: 2					
328	NUMBER OF PARAMETERS: 3					
329	TYPE OF MODEL: PROBIT					
330	THERE ARE TRANSFORMATIONS					
331	INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY					
332	ALL RAW DATA WILL BE PRINTED					
333	NATURAL RESPONSE IS NOT A PARAMETER					
334	MULTIPLE SUBJECTS PER DATA CARD					
335	TRANSFORMATIONS: 2L2=3L3=					
336	PARAMETER LABELS					
337	1 CONSTANT					
338	2 LOG(C)					
339	3 LOG(A)					
340	FORMAT: (1X,1X,2F5.0,2I4)					
341	INPUT DATA					
342	1 5.	2.9	27	1		
341a	1 10.	2.9	29	15		
342a	1 20.	2.9	30	27		
343	1 40.	2.9	28	28		
344	1 5.	5.7	29	4		
345	1 10.	5.7	29	19		
346	1 20.	5.7	27	26		
347	1 40.	5.7	30	30		
348	1 5.	10.8	30	6		
349	1 10.	10.8	24	15		
350	1 20.	10.8	31	31		
351	1 40.	10.8	19	19		
352	2 5.	2.9	29	3		
353	2 10.	2.9	30	10		
354	2 20.	2.9	29	24		
355	2 40.	2.9	29	29		
356	2 5.	5.7	27	4		
357	2 10.	5.7	28	14		
358	2 20.	5.7	28	27		
359	2 40.	5.7	29	29		
360	2 5.	10.8	28	8		
361	2 10.	10.8	28	17		
362	2 20.	10.8	28	26		
363	2 40.	10.8	17	17		
364	TRANSFORMED DATA					
365	1.	1	1.00000	.698970	.462398	
366	2.	0	1.00000	.698970	.462398	
367	3.	0	1.00000	.698970	.462398	
368	4.	0	1.00000	.698970	.462398	
369	5.	0	1.00000	.698970	.462398	
370	6.	0	1.00000	.698970	.462398	
371	7.	0	1.00000	.698970	.462398	
372	8.	0	1.00000	.698970	.462398	
373	9.	0	1.00000	.698970	.462398	
374	10.	0	1.00000	.698970	.462398	
375	11.	0	1.00000	.698970	.462398	
377	12.	0	1.00000	.698970	.462398	
378	PYR SF NONAT EQUALITY	POL02	26 AUG 80	14:07:05	19	39
379	613.	0	1.00000	1.00000	1.03342	
380	614.	0	1.00000	1.00000	1.03342	
381	615.	0	1.00000	1.00000	1.03342	
382	616.	0	1.00000	1.00000	1.03342	
383	617.	0	1.00000	1.00000	1.03342	
384	618.	0	1.00000	1.00000	1.03342	
385	619.	1	1.00000	1.30103	1.03342	
386	620.	1	1.00000	1.30103	1.03342	
387	621.	1	1.00000	1.30103	1.03342	
388	622.	1	1.00000	1.30103	1.03342	
389	623.	1	1.00000	1.30103	1.03342	
390	624.	1	1.00000	1.30103	1.03342	
391	625.	1	1.00000	1.30103	1.03342	
392	626.	1	1.00000	1.30103	1.03342	
393	627.	1	1.00000	1.30103	1.03342	
394	628.	1	1.00000	1.30103	1.03342	
395	629.	1	1.00000	1.30103	1.03342	
396	630.	1	1.00000	1.30103	1.03342	
397	631.	1	1.00000	1.30103	1.03342	
398	632.	1	1.00000	1.30103	1.03342	
399	633.	1	1.00000	1.30103	1.03342	
400	634.	1	1.00000	1.30103	1.03342	
401	635.	1	1.00000	1.30103	1.03342	

Figure 10-4—Continued

```

402  636.  1  1.00000  1.30103  1.03342
403  637.  1  1.00000  1.30103  1.03342
404  638.  1  1.00000  1.30103  1.03342
405  639.  1  1.00000  1.30103  1.03342
406  640.  1  1.00000  1.30103  1.03342
407  641.  1  1.00000  1.30103  1.03342
408  642.  1  1.00000  1.30103  1.03342
409  643.  1  1.00000  1.30103  1.03342
410  644.  1  1.00000  1.30103  1.03342
411  645.  0  1.00000  1.30103  1.03342
412  646.  0  1.00000  1.30103  1.03342
413  647.  1  1.00000  1.60206  1.03342
414  648.  1  1.00000  1.60206  1.03342
415  649.  1  1.00000  1.60206  1.03342
416  650.  1  1.00000  1.60206  1.03342
417  651.  1  1.00000  1.60206  1.03342
418  652.  1  1.00000  1.60206  1.03342
419  653.  1  1.00000  1.60206  1.03342
420  654.  1  1.00000  1.60206  1.03342
421  655.  1  1.00000  1.60206  1.03342
422  656.  1  1.00000  1.60206  1.03342
423  657.  1  1.00000  1.60206  1.03342
424  658.  1  1.00000  1.60206  1.03342
425  659.  1  1.00000  1.60206  1.03342
426  660.  1  1.00000  1.60206  1.03342
427  661.  1  1.00000  1.60206  1.03342
428  662.  1  1.00000  1.60206  1.03342
429  663.  1  1.00000  1.60206  1.03342

430  663 OBSERVATIONS + 0 CONTROLS = 663 TOTAL

431  INITIAL ESTIMATES OF THE PARAMETERS
432      -4.544105 CONSTANT          3.947060 LOG(C)          1.018223 LOG(W)

433  STARTING LIKELIHOOD = -226.9796

434  6 LOGIT EVALUATIONS

```

```

435  PYR SF NONAT EQUALITY                                POL02  26 AUG 80  14:07:13  22  40

```

```

436  6 PROBIT EVALUATIONS

```

```

437  MAXIMUM LOG LIKELIHOOD = -225.8631

```

	PARAMETER	STD ERROR	T RATIO
438	CONSTANT	-5.03989	.408145
439	LOG(C)	4.37818	.298517
440	LOG(W)	1.12486	.293838

```

442  COVARIANCE MATRIX

```

	CONSTANT	LOG(C)	LOG(W)
443	CONSTANT	.1665827	-.1021096
444	LOG(C)	-.1021096	.8911216-01
445	LOG(W)	-.8328212-01	.1960563-01

```

447  PREDICTION SUCCESS TABLE

```

ACTUAL OUTCOME	PREDICTED OUTCOME	
	ALIVE	DEAD
448		
449		
450		
451		
452	ALIVE	178
453	DEAD	51
454		
455	PERCENT CORRECT	78
456	OVERALL PERCENT CORRECT: 84	87

```

457  TESTING HYPOTHESIS THAT THE 3-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
458  LIKELIHOOD RATIO TEST: CHI-SQUARED = 467.3855  DF=3  TAIL PROBABILITY = .000

```

Figure 10-5

```

459  PYRETHRIN SPRAY OR FILM WITH NATURAL RESPONSE                                POL02          28 AUG 80  09:33:01  10  4

460  CONTROL CARD: 4,2,0,3,1,0,1,0,1,2,0,0

461  NUMBER OF VARIABLES: 4
462  NUMBER OF VARIABLES ON INPUT CARDS: 2
463  NUMBER OF PARAMETERS: 5
464  TYPE OF MODEL: PROBIT
465  THERE ARE TRANSFORMATIONS

```

Figure 10-5—Continued

466 A CONTROL GROUP IS INDICATED BY VARIABLE 3 BEING ZERO
 467 INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY
 468 ALL RAW DATA WILL BE PRINTED
 469 NATURAL RESPONSE IS A PARAMETER
 470 MULTIPLE SUBJECTS PER DATA CARD
 471 NUMBER OF PARALLEL GROUPS: 2

472 TRANSFORMATIONS: 3L3=4L4=

473 PARAMETER LABELS

474 1 SPRAY
 475 2 FILM
 476 3 LOG(C)
 477 4 LOG(W)
 478 5 NATURAL

479 FORMAT: (I1,1X,2F5.0,2I4)

480 INPUT DATA

481	1 5.	2.9	27	1
482	1 10.	2.9	29	15
483	1 20.	2.9	30	27
484	1 40.	2.9	28	28
485	1 5.	5.7	29	4
486	1 10.	5.7	29	19
487	1 20.	5.7	27	26
488	1 40.	5.7	30	30
489	1 5.	10.8	30	6
490	1 10.	10.8	29	15
491	1 20.	10.8	31	31
492	1 40.	10.8	19	19
493	2 5.	2.9	29	3
494	2 10.	2.9	30	10
495	2 20.	2.9	29	24
496	2 40.	2.9	29	29
497	2 5.	5.7	27	4
498	2 10.	5.7	28	14
499	2 20.	5.7	28	27
500	2 40.	5.7	29	24
501	2 5.	10.8	28	8
502	2 10.	10.8	28	17
503	2 20.	10.8	28	26
504	2 40.	10.8	17	17
505	0 0	0	311	12

506 TRANSFORMED DATA

507	1.	1	1.000000	.000000	.698970	.462398
508	2.	0	1.000000	.000000	.698970	.462398
509	3.	0	1.000000	.000000	.698970	.462398
510	4.	0	1.000000	.000000	.698970	.462398
511	5.	0	1.000000	.000000	.698970	.462398
512	6.	0	1.000000	.000000	.698970	.462398
513	7.	0	1.000000	.000000	.698970	.462398

514 PYRETHRUM SPRAY OR FILM WITH NATURAL RESPONSE

POL02 28 AUG 80 09:33:02 12 15

515	608.	0	.000000	1.00000	1.00000	1.03342
516	609.	0	.000000	1.00000	1.00000	1.03342
517	610.	0	.000000	1.00000	1.00000	1.03342
518	611.	0	.000000	1.00000	1.00000	1.03342
519	612.	0	.000000	1.00000	1.00000	1.03342
520	613.	0	.000000	1.00000	1.00000	1.03342
521	614.	0	.000000	1.00000	1.00000	1.03342
522	615.	0	.000000	1.00000	1.00000	1.03342
523	616.	0	.000000	1.00000	1.00000	1.03342
524	617.	0	.000000	1.00000	1.00000	1.03342
525	618.	0	.000000	1.00000	1.00000	1.03342
526	619.	1	.000000	1.00000	1.30103	1.03342
527	620.	1	.000000	1.00000	1.30103	1.03342
528	621.	1	.000000	1.00000	1.30103	1.03342
529	622.	1	.000000	1.00000	1.30103	1.03342
530	623.	1	.000000	1.00000	1.30103	1.03342
531	624.	1	.000000	1.00000	1.30103	1.03342
532	625.	1	.000000	1.00000	1.30103	1.03342
533	626.	1	.000000	1.00000	1.30103	1.03342
534	627.	1	.000000	1.00000	1.30103	1.03342
535	628.	1	.000000	1.00000	1.30103	1.03342
536	629.	1	.000000	1.00000	1.30103	1.03342
537	630.	1	.000000	1.00000	1.30103	1.03342
538	631.	1	.000000	1.00000	1.30103	1.03342
539	632.	1	.000000	1.00000	1.30103	1.03342
540	633.	1	.000000	1.00000	1.30103	1.03342
541	634.	1	.000000	1.00000	1.30103	1.03342
542	635.	1	.000000	1.00000	1.30103	1.03342
543	636.	1	.000000	1.00000	1.30103	1.03342
544	637.	1	.000000	1.00000	1.30103	1.03342
545	638.	1	.000000	1.00000	1.30103	1.03342
546	639.	1	.000000	1.00000	1.30103	1.03342
547	640.	1	.000000	1.00000	1.30103	1.03342

Figure 10-5—Continued

548	641.	1	.000000	1.00000	1.30103	1.03342
549	642.	1	.000000	1.00000	1.30103	1.03342
550	643.	1	.000000	1.00000	1.30103	1.03342
551	644.	1	.000000	1.00000	1.30103	1.03342
552	645.	0	.000000	1.00000	1.30103	1.03342
553	646.	0	.000000	1.00000	1.30103	1.03342
554	647.	1	.000000	1.00000	1.60206	1.03342
555	648.	1	.000000	1.00000	1.60206	1.03342
556	649.	1	.000000	1.00000	1.60206	1.03342
557	650.	1	.000000	1.00000	1.60206	1.03342
558	651.	1	.000000	1.00000	1.60206	1.03342
559	652.	1	.000000	1.00000	1.60206	1.03342
560	653.	1	.000000	1.00000	1.60206	1.03342
561	654.	1	.000000	1.00000	1.60206	1.03342
562	655.	1	.000000	1.00000	1.60206	1.03342
563	656.	1	.000000	1.00000	1.60206	1.03342
564	657.	1	.000000	1.00000	1.60206	1.03342
565	658.	1	.000000	1.00000	1.60206	1.03342
566	659.	1	.000000	1.00000	1.60206	1.03342
567	660.	1	.000000	1.00000	1.60206	1.03342
568	661.	1	.000000	1.00000	1.60206	1.03342
569	662.	1	.000000	1.00000	1.60206	1.03342
570	663.	1	.000000	1.00000	1.60206	1.03342

571 663 OBSERVATIONS + 311 CONTROLS = 974 TOTAL
 572 MORTALITY IN CONTROLS: .039

573 PYRETHRUM SPRAY OR FILM WITH NATURAL RESPONSE

POLO2 28 AUG 80 09:33:04 14 16

574 INITIAL ESTIMATES OF THE PARAMETERS

575 -4.101413 SPRAY .8847998 LOG(W)
 576 -4.209135 FILM .3858521-01 NATURAL
 577 3.631875 LOG(C)

578 STARTING LIKELIHOOD = -281.2558

579 10 LOGIT EVALUATIONS

580 5 PROBIT EVALUATIONS

581 MAXIMUM LOG LIKELIHOOD = -275.5481

	PARAMETER	STD ERROR	T RATIO	
582				
583	SPRAY	-5.41446	.504502	-10.7323
584	FILM	-5.61829	.525942	-10.6823
585	LOG(C)	4.70390	.365933	12.8546
586	LOG(W)	1.22311	.327875	3.73041
587	NATURAL	.345908-01	.110382-01	3.58671

588 COVARIANCE MATRIX

	SPRAY	FILM	LOG(C)	LOG(W)	NATURAL	
589						
590	SPRAY	.2545221	.2543949	-.1591772	-.1135423	-.1383305-02
591	FILM	.2543949	.2766148	-.1676289	-.1163738	-.1537166-02
592	LOG(C)	-.1591772	-.1676289	.1339068	.3393508-01	.1019743-02
593	LOG(W)	-.1135423	-.1163738	.3393508-01	.1075022	.2416572-03
594	NATURAL	-.1383305-02	-.1537166-02	.1019743-02	.2416572-03	.1218416-03

595 PREDICTION SUCCESS TABLE

ACTUAL OUTCOME	PREDICTED OUTCOME		
	ALIVE	DEAD	
596			
597			
598			
599			
600	ALIVE	178	56
601	DEAD	51	378
602			
603	PERCENT CORRECT	78	87
604	OVERALL PERCENT CORRECT: 84		

605 TESTING HYPOTHESIS THAT THE 4-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
 606 LIKELIHOOD RATIO TEST: CHI-SQUARED = 455.2636 DF=4 TAIL PROBABILITY = .000

4.1.6 Hypotheses Testing

The maximized log likelihood values needed to test the hypotheses outlined in section 4.1.2 are:

Model	L	Source
[1]	-101.3851	fig. 10-1, line 72
[2]	-122.4908	fig. 10-2, line 162
[3]	-225.1918	fig. 10-3, line 300
[4]	-225.8631	fig. 10-4, line 437

The LR tests of the three hypotheses are the following:

(1) Hypothesis H(P) of parallelism.

$$\begin{aligned} L(\Omega) &= -101.3851 + -122.4908 \\ &= -223.8759, \\ L(\omega) &= -225.1918, \\ LR &= 2[L(\Omega) - L(\omega)] = 2[-223.8759 + 225.1918] \\ &= 2[1.3159] = 2.6318. \end{aligned}$$

The hypothesis H(P) is accepted, at significance level $\alpha = 0.05$ if

$$LR \leq \chi^2_{.05}(2) = 5.99.$$

Since $2.6318 < 5.99$, we accept H(P).

(2) Hypothesis H(E|P) of equality given parallelism.

$$\begin{aligned} L(\Omega) &= -225.1918, \\ L(\omega) &= -225.8631, \\ LR &= 2[L(\Omega) - L(\omega)] = 2[-225.1918 + 225.8631] \\ &= 2[0.6713] = 1.3426. \end{aligned}$$

The hypothesis H(E|P) is accepted at level $\alpha = 0.05$ if

$$LR \leq \chi^2_{.05}(1) = 3.84.$$

Since $1.3426 < 3.84$, we accept H(E|P).

(3) Hypothesis H(E) of equality.

$$\begin{aligned} L(\Omega) &= -223.8759, \\ L(\omega) &= -225.8631, \\ LR &= 2[L(\Omega) - L(\omega)] = 2[-223.8759 + 225.8631] \\ &= 2[1.9872] = 3.9744. \end{aligned}$$

The hypothesis H(E) is accepted at level $\alpha = 0.05$ if

$$LR \leq \chi^2_{.05}(3) = 7.81.$$

We also accept this hypothesis since $3.9744 < 7.81$.

4.1.7 Comparison with Published Calculations

The analyses of models [1]-[4] cannot be compared directly with those described by Finney. The parameter values for model [5] confirm those of Finney's equations (8.27) and (8.28) (Finney 1971, p. 169).

4.2 Vaso-Constriction

Finney (1971, p. 183-190) describes a series of measurements of the volume of air inspired by human subjects, their rate of inspiration, and whether or not a vaso-constriction reflex occurred in the skin of their fingers. These experiments were reported originally by

Gilliatt (1947). A feature of this example is that the data are ungrouped. We use the example to illustrate the analysis of ungrouped data, the likelihood ratio test for equal regression coefficients, and the use of transformations.

4.2.1 Models, Hypothesis, and Analyses Required

The model expressing the probability of the vaso-constriction reflex is

$$Y = \alpha + \beta_1 x_1 + \beta_2 x_2 \quad [1]$$

where y is the probit or logit of the probability, α is the constant term, x_1 is the logarithm volume of air inspired in liters, x_2 is the logarithm of rate of inspiration in liters per second, β_1 is the regression coefficient for volume, and β_2 is the regression coefficient for rate.

The hypothesis is $H: \beta_1 = \beta_2$ which states that the regression coefficients for rate and volume of air inspired are the same. The unrestricted model is [1] and the restricted model is

$$Y = \alpha + \beta(x_1 + x_2) = \alpha + \beta_x \quad [2]$$

The required maximum log likelihoods are $L(\Omega)$ and $L(\omega)$.

- (i) $L(\Omega)$: ML estimation of [1].
- (ii) $L(\omega)$: ML estimation of [2].

When the hypothesis H is true, asymptotically,

$$LR = 2[L(\Omega) - L(\omega)] \sim \chi^2(1)$$

so that the hypothesis H is accepted at the α level of significance if

$$LR \leq \chi^2_{\alpha}(1).$$

4.2.2 Input

The input for analyses of the two required models consists of 97 cards (fig. 11). After the starter cards (fig. 11-A), program cards specify the analysis of model [1] (fig. 11-B-I). The title card (card 6) cites the source of the data; the control card (card 7) specifies three regression coefficients, two variables to be read from each data card, the probit model to be used, none of the explanatory variables contains a control group, transformations will be used, starting values of the parameters will be calculated automatically, data will be printed back, natural response is not a parameter, there is one subject per data card, there are no parallel groups, a restricted model will not be computed, and the program will select a suitable number of iterations in search of an ML estimate. The transformation card (card 8) defines y_2 as the logarithm of x_2 , and y_3 as the logarithm of x_3 . The three parameters (card 9) are labeled constant (x_1), volume (x_2), and rate (x_3). The format statement (card 10) instructs the program to read a 5-column field including a decimal point (volume), and a 6-column field including a decimal point (rate), and finally, a single column of integers (1=constricted, 0=not constricted). The data, consisting of 39 individual records,

B-1

```

1 @RUN,8 IRV,143-8801,,3,100 POLO2
2 @ASG,T X.,T U20009
3 @COPY,G X.
4 @FREE X.
5 @XQT POLO2
6 =VASO-CONSTRICTION EXAMPLE FROM FINNEY,PROBIT ANALYSIS,PP.183-190
7 3,2,0,0,1,0,1,1,0,0,0,0
8 2L2=3L3=
9 CONSTANT VOLUME RATE
10 (F5.0,F6.0,I1)
11 3.7 0.825 1
12 3.5 1.09 1
13 1.25 2.5 1
14 0.75 1.5 1
15 0.8 3.2 1
16 0.7 3.5 1
17 0.6 0.75 0
18 1.1 1.7 0
19 0.9 0.75 0
20 0.9 0.45 0
21 0.8 0.57 0
22 0.55 2.75 0
23 0.6 3.0 0
24 1.4 2.33 1
25 0.75 3.75 1
26 2.3 1.64 1
27 3.2 1.6 1
28 0.85 1.415 1
29 1.7 1.06 0
30 1.8 1.8 1
31 0.4 2.0 0
32 0.95 1.36 0
33 1.35 1.35 0
34 1.5 1.36 0
35 1.6 1.78 1
36 0.6 1.5 0
37 1.8 1.5 1
38 0.95 1.9 0
39 1.9 0.95 1
40 1.6 0.4 0
41 2.7 0.75 1
42 2.35 0.3 0
43 1.1 1.83 0
44 1.1 2.2 1
45 1.2 2.0 1
46 0.8 3.33 1
47 0.95 1.9 0
48 0.75 1.9 0
49 1.3 1.625 1
50 END
51 =VASOCON EQUAL SLOPES
52 2,2,0,0,1,0,1,1,0,0,0,0
53 2L3L+2=
54 CONSTANT COMBINE
55 (F5.0,F6.0,I1)
56 3.7 0.825 1
57 1.3 1.625 1
58 END

```

C-2

```

1 3.7 0.825 1
2 3.5 1.09 1
3 1.25 2.5 1
4 0.75 1.5 1
5 0.8 3.2 1
6 0.7 3.5 1
7 0.6 0.75 0
8 1.1 1.7 0
9 0.9 0.75 0
10 0.9 0.45 0
11 0.8 0.57 0
12 0.55 2.75 0
13 0.6 3.0 0
14 1.4 2.33 1
15 0.75 3.75 1
16 2.3 1.64 1
17 3.2 1.6 1
18 0.85 1.415 1
19 1.7 1.06 0
20 1.8 1.8 1
21 0.4 2.0 0
22 0.95 1.36 0
23 1.35 1.35 0
24 1.5 1.36 0
25 1.6 1.78 1
26 0.6 1.5 0
27 1.8 1.5 1
28 0.95 1.9 0
29 1.9 0.95 1
30 1.6 0.4 0
31 2.7 0.75 1
32 2.35 0.3 0
33 1.1 1.83 0
34 1.1 2.2 1
35 1.2 2.0 1
36 0.8 3.33 1
37 0.95 1.9 0
38 0.75 1.9 0
39 1.3 1.625 1
40 END
41 =VASOCON EQUAL SLOPES
42 2,2,0,0,1,0,1,1,0,0,0,0
43 2L3L+2=
44 CONSTANT COMBINE
45 (F5.0,F6.0,I1)
46 3.7 0.825 1
47 1.3 1.625 1
48 END

```

follows (fig. 11-C-1). After the END card (card 51), the input for analysis of model [2] follows.

4.2.3 Output

rate. The summary of total observations concludes the descriptive portion of each printout (fig. 12, lines 99 and 225).

Initial parameter estimates, starting ML values, and iteration statements (fig. 12, lines 100-104; lines 226-230) begin the statistical portion of each printout. The final ML estimate follows (fig. 12, lines 105 and 231); parameter values, their standard errors, and t-ratios (fig. 12, lines 106-109; lines 232-234) are printed next. The covariance matrix and prediction success table follow (fig. 12, lines 111-125; lines 236-249). Note that the program's automatic dead-alive category labels are not appropriate for this experiment; labels such as constricted and not constricted would be more appropriate. The LR test for significance of the model coefficients ends each analysis (fig. 12, lines 126-127; lines 250-251).

The t-ratios of the parameters for both models indicate that each parameter is significant in the regression. The

OPC values indicate that each model is a good predictor of observed results, and the LR tests indicate that both regressions are highly significant ($\alpha = 0.05$).

4.2.4 Hypothesis Testing

The maximum log likelihoods needed to test the hypothesis $H: \beta_1 = \beta_2$ are:

Model	L	Source
[1]	-14.6608	fig. 12, line 105
[2]	-14.7746	fig. 12, line 231

In this example

$$\begin{aligned} L(\Omega) &= -14.6608, \\ L(\omega) &= -14.7746, \\ LR &= 2[L(\Omega) - L(\omega)] = 2[0.1138] = 0.2276. \end{aligned}$$

For a test at the 0.05 significance levels, the χ^2 critical value is 3.84. Hence we accept the hypothesis H.

Figure 12

1	VASO-CONSTRICTION,EXAMPLE FROM FINNEY,PROBIT ANALYSIS,PP.183-190	POL02	28 AUG 80	09:33:07	17	17
2	CONTROL CARD: 3,2,0,0,1,0,1,1,0,0,0,0					
3	NUMBER OF VARIABLES: 3					
4	NUMBER OF VARIABLES ON INPUT CARDS: 2					
5	NUMBER OF PARAMETERS: 3					
6	TYPE OF MODEL: PROBIT					
7	THERE ARE TRANSFORMATIONS					
8	INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY					
9	ALL RAW DATA WILL BE PRINTED					
10	NATURAL RESPONSE IS NOT A PARAMETER					
11	ONE SUBJECT PER DATA CARD					
12	TRANSFORMATIONS: 2L2=3L3=					
13	PARAMETER LABELS					
14	1 CONSTANT					
15	2 VOLUME					
16	3 RATE					
17	FORMAT: (F5.0,F6.0,11)					
18	INPUT DATA					
19	3.7 0.425 1					
20	3.5 1.09 1					
21	1.25 2.5 1					
22	0.75 1.5 1					
23	0.8 3.2 1					
24	0.7 3.5 1					
25	0.6 0.75 0					
26	1.1 1.7 0					
27	0.9 0.75 0					
28	0.9 0.45 0					
29	0.8 0.57 0					
30	0.55 2.75 0					
31	0.6 5.0 0					
32	1.4 2.53 1					
33	0.75 5.75 1					
34	2.3 1.64 1					
35	3.2 1.6 1					
36	0.85 1.415 1					
37	1.7 1.06 0					
38	1.8 1.8 1					
39	0.4 2.0 0					
40	0.95 1.36 0					
41	1.45 1.55 0					
42	1.5 1.36 0					
43	1.6 1.78 1					

Figure 12—Continued

```

44 0.4 1.5 0
45 1.8 1.5 1
46 0.95 1.9 0
47 1.9 0.95 1
48 1.6 0.4 0
49 2.7 0.75 1
50 2.35 0.3 0
51 1.1 1.83 0
52 1.1 2.2 1
53 1.7 2.0 1
54 0.8 3.33 1
55 0.95 1.9 0
56 0.75 1.9 0

```

```

57 VASO-CONSTRICTION,EXAMPLE FROM FINNEY,PROBIT ANALYSIS,PP.183-190 POL02 28 AUG 80 09:33:08 17 18

```

```

58 1.3 1.625 1

```

59 TRANSFORMED DATA

60	1.	1	1.00000	.568202	-.835460-01
61	2.	1	1.00000	.544068	.374265-01
62	3.	1	1.00000	.969100-01	.397940
63	4.	1	1.00000	-.124939	.176091
64	5.	1	1.00000	-.969100-01	.505150
65	6.	1	1.00000	-.154902	.544068
66	7.	0	1.00000	-.221849	-.124939
67	8.	0	1.00000	.413927-01	.230449
68	9.	0	1.00000	-.457575-01	-.124939
69	10.	0	1.00000	-.457575-01	-.346787
70	11.	0	1.00000	-.969100-01	-.244125
71	12.	0	1.00000	-.259637	.439333
72	13.	0	1.00000	-.221849	.477121
73	14.	1	1.00000	.146128	.367356
74	15.	1	1.00000	-.124939	.574031
75	16.	1	1.00000	.361728	.214844
76	17.	1	1.00000	.505150	.204120
77	18.	1	1.00000	-.705811-01	.150756
78	19.	0	1.00000	.230449	.253059-01
79	20.	1	1.00000	.255273	.255273
80	21.	0	1.00000	-.397940	.301030
81	22.	0	1.00000	-.222764-01	.133539
82	23.	0	1.00000	.130334	.130334
83	24.	0	1.00000	.176091	.133539
84	25.	1	1.00000	.204120	.250420
85	26.	0	1.00000	-.221849	.176091
86	27.	1	1.00000	.255273	.176091
87	28.	0	1.00000	-.222764-01	.278754
88	29.	1	1.00000	.278754	-.222764-01
89	30.	0	1.00000	.204120	-.397940
90	31.	1	1.00000	.431364	-.124939
91	32.	0	1.00000	.571068	-.522879
92	33.	0	1.00000	.413927-01	.262451
93	34.	1	1.00000	.413927-01	.342423
94	35.	1	1.00000	.791812-01	.301030
95	36.	1	1.00000	-.969100-01	.522444
96	37.	0	1.00000	-.222764-01	.278754
97	38.	0	1.00000	-.124939	.278754
98	39.	1	1.00000	.113943	.210853

99 39 OBSERVATIONS + 0 CONTROLS = 39 TOTAL

100 INITIAL ESTIMATES OF THE PARAMETERS

101 -.2624056 CONSTANT 1.637666 VOLUME 1.240651 RATE

102 STARTING LIKELIHOOD = -20.5856

103 14 LOGIT EVALUATIONS

104 8 PROBIT EVALUATIONS

105 MAXIMUM LOG LIKELIHOOD = -14.6608

106	PARAMETER	STD ERROR	T RATIO
107	CONSTANT -1.53306	.611100	-2.50869
108	VOLUME 6.63059	2.08127	3.18584
109	RATE 5.88569	2.06139	2.85521

```

110 VASO-CONSTRICTION,EXAMPLE FROM FINNEY,PROBIT ANALYSIS,PP.183-190 POL02 28 AUG 80 09:33:08 18 19

```

111 COVARIANCE MATRIX

112	CONSTANT	VOLUME	RATE
113	CONSTANT .3734435	-.9039205	-1.128364
114	VOLUME -.9039205	4.331670	3.066827
115	RATE -1.128364	3.066827	4.249311

Figure 12—Continued

116	PREDICTION SUCCESS TABLE		
117		PREDICTED OUTCOME	
118		-----	
119	ACTUAL OUTCOME	ALIVE	DEAD
120		-----	
121	ALIVE	14	5
122	DEAD	2	18
123		-----	
124	PERCENT CORRECT	88	78
125	OVERALL PERCENT CORRECT: 82		

126 TESTING HYPOTHESIS THAT THE 3-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
 127 LIKELIHOOD RATIO TEST: CHI-SQUARED = 24.7440 DF=3 TAIL PROBABILITY = .000

128 VASOCON EQUAL SLOPES POL02 28 AUG 80 09:33:08 18 20

129 CONTROL CARD: 2,2,0,0,1,0,1,1,0,0,0,0

130 NUMBER OF VARIABLES: 2

131 NUMBER OF VARIABLES ON INPUT CARDS: 2

132 NUMBER OF PARAMETERS: 2

133 TYPE OF MODEL: PROBIT

134 THERE ARE TRANSFORMATIONS

135 INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY

136 ALL RAW DATA WILL BE PRINTED

137 NATURAL RESPONSE IS NOT A PARAMETER

138 ONE SUBJECT PER DATA CARD

139 TRANSFORMATIONS: 2L3L+2=

140 PARAMETER LABELS

141 1 CONSTANT

142 2 COMBINE

143 FORMAT: (F5.0,F6.0,I1)

144 INPUT DATA

145 3.7 0.825 1

146 3.5 1.09 1

147 1.25 2.5 1

148 0.75 1.5 1

149 0.8 3.2 1

150 0.7 3.5 1

151 0.6 0.75 0

152 1.1 1.7 0

153 0.9 0.75 0

154 0.9 0.45 0

155 0.8 0.57 0

156 0.55 2.75 0

157 0.6 3.0 0

158 1.4 2.43 1

159 0.75 3.75 1

160 2.3 1.64 1

161 3.2 1.6 1

162 0.85 1.415 1

163 1.7 1.06 0

164 1.8 1.8 1

165 0.4 2.0 0

166 0.95 1.36 0

167 1.35 1.35 0

168 1.5 1.36 0

169 1.6 1.74 1

170 0.6 1.5 0

171 1.8 1.5 1

172 0.95 1.9 0

173 1.9 0.95 1

174 1.6 0.4 0

175 2.7 0.75 1

176 2.35 0.5 0

177 1.1 1.83 0

178 1.1 2.2 1

179 1.2 2.0 1

180 0.8 3.33 1

181 0.95 1.9 0

182 0.75 1.9 0

183 1.3 1.625 1

184 VASOCON EQUAL SLOPES POL02 28 AUG 80 09:33:08 18 21

185 TRANSFORMED DATA

186 1. 1 1.00000 .484656

187 2. 1 1.00000 .581495

188 3. 1 1.00000 .494850

Figure 12—Continued

```

189      4.  1  1.00000  .511525-01
190      5.  1  1.00000  .408240
191      6.  1  1.00000  .389166
192      7.  0  1.00000  -.345787
193      8.  0  1.00000  .271842
194      9.  0  1.00000  -.170696
195     10.  0  1.00000  -.392545
196     11.  0  1.00000  -.341035
197     12.  0  1.00000  .179695
198     13.  0  1.00000  .255273
199     14.  1  1.00000  .513484
200     15.  1  1.00000  .449093
201     16.  1  1.00000  .576572
202     17.  1  1.00000  .709270
203     18.  1  1.00000  .801754-01
204     19.  0  1.00000  .255755
205     20.  1  1.00000  .510545
206     21.  0  1.00000  -.969100-01
207     22.  0  1.00000  .111263
208     23.  0  1.00000  .260668
209     24.  0  1.00000  .309630
210     25.  1  1.00000  .454540
211     26.  0  1.00000  -.457575-01
212     27.  1  1.00000  .431364
213     28.  0  1.00000  .256477
214     29.  1  1.00000  .256477
215     30.  0  1.00000  -.193820
216     31.  1  1.00000  .306425
217     32.  0  1.00000  -.151811
218     33.  0  1.00000  .303844
219     34.  1  1.00000  .483815
220     35.  1  1.00000  .380211
221     36.  1  1.00000  .425534
222     37.  0  1.00000  .256477
223     38.  0  1.00000  .153815
224     39.  1  1.00000  .424797

```

225 39 OBSERVATIONS + 0 CONTROLS = 39 TOTAL

226 INITIAL ESTIMATES OF THE PARAMETERS

227 -.740226 CONSTANT 3.077000 COMBINE

228 STARTING LIKELIHOOD = -14.7708

229 9 LOGIT EVALUATIONS

230 5 PROBIT EVALUATIONS

231 MAXIMUM LOG LIKELIHOOD = -14.7746

	PARAMETER	STD ERROR	T RATIO
232 CONSTANT	-1.61648	.602371	-2.68353
233 COMBINE	6.30556	1.93851	3.25279

235 VASOCON EQUAL SLOPES

POL02

28 AUG 80 09:33:09 19 22

236 COVARIANCE MATRIX

	CONSTANT	COMBINE
237 CONSTANT	.3628504	-1.049731
238 COMBINE	-1.049731	3.757826

240 PREDICTION SUCCESS TABLE

ACTUAL OUTCOME	PREDICTED OUTCOME	
	ALIVE	DEAD
241 ALIVE	13	6
242 DEAD	2	18
243 PERCENT CORRECT	87	75
244 OVERALL PERCENT CORRECT: 79		

250 TESTING HYPOTHESIS THAT THE 2-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME

251 LIKELIHOOD RATIO TEST: CHI-SQUARED = 24.5162 DF=2 TAIL PROBABILITY = .000

4.3 Body Weight as a Variable: Higher Order Terms

Robertson and others (1981) described a series of experiments designed to test the hypothesis that the response of an insect (*Choristoneura occidentalis* Freeman) is proportional to its body weight. Three chemicals, including mexacarbate, were used. We use data for tests with mexacarbate to illustrate the use of individual data to test the significance of a higher order term in a polynomial model. This example demonstrates the use of the restricted model option of POLO2 (see *section 2.3*).

Briefly, each insect in this experiment was selected at random from a laboratory colony, weighed, and treated with 1 μ l of mexacarbate dissolved in acetone. Mortality was tallied after 7 days. Individual records for 253 insects were kept. Because a printback of all the data is too voluminous for this report, we use the program option of IECHO=0 (see *section 2.3*).

4.3.1 Models and Hypothesis

The polynomial model is

$$y = \delta_0 + \delta_1 x_1 + \delta_2 x_2 + \delta_3 z$$

where y is the probit or logit of response, x_1 is the logarithm of dose (in μg), x_2 is the logarithm of body weight (in mg), and z is the square of the logarithm of body weight ($z=x_2^2$). The regression coefficients are δ_0 for the constant, δ_1 for log dose, δ_2 for log weight, and δ_3 for the square of log weight.

The hypothesis is $H: \delta_3=0$, that is, the coefficient of the higher order term in weight equals zero. The unrestricted model is [1] and the restricted model is

$$Y = \delta_0 + \delta_1 X_1 + \delta_2 X_2 \quad [2]$$

The required maximum log likelihoods are $L(\Omega)$ and $L(\omega)$.

- (i) $L(\Omega)$: ML estimation of [1].
- (ii) $L(\omega)$: ML estimation of [2].

When H is true, asymptotically,

$$\text{LR} = 2[\text{L}(\Omega) - \text{L}(\omega)] \sim \chi^2(1),$$

so that H is accepted at level α if

$$\text{LR} \leq \chi^2 (1).$$

The program will automatically conduct an LR test of $H: \delta_3=0$.

4.3.2 Input

The input for this analysis consists of 263 cards (*fig. 13*). The program cards (*fig. 13-B*) follow the usual starter cards (*fig. 13-A*). Following the title (card 6), the control card specifies four regression coefficients, two variables to be read from each data card, the probit model is to be used, the second parameter (log (D)) contains a control group, transformations will be used, starting values will be calculated automatically, data printback will be suppressed, natural response is a parameter, there is one subject per data card, there are no parallel groups, a restricted model retaining three variables will be computed, and the program will choose the number of iterations for M L estimates. The transformations defined (card 8) are: y_2 equals the logarithm of x_2 , y_3 equals the logarithm of x_3 , and y_4 equals the square of the logarithm of x_3 . Parameters are x_1 ="constant," x_2 ="logarithm of dose," x_3 ="logarithm of weight," x_4 ="($\log$ arithm of weight)²," and x_5 ="natural response." Suggestive labels appear on card 9.

The germane information on each data card (card 11-263, *fig. 13-C*) is listed in columns 12-14 (dose), 17-19 (weight), and 24 (dead or alive). The format statement (card 10), therefore, instructs the program to skip the first 10 columns, read a 4-column field with two digits to the right of the decimal point, read a 5-column field with one digit to the right of the decimal point, then go to column 24 to read an integer. No END card is needed after the data assuming that no other analysis follows.

[illegible]

Figure 13

4.3.3 Output

The title is repeated at the top of each page of the printout (*fig. 14*, lines 1, 56, and 101). The initial portions of the analysis demonstrate the same features noted in previous examples: the control card listing (*fig. 14*, line 2) is followed by specification statements (*fig. 14*, lines 3-12), the transformation statement (*fig. 14*, line 13), parameter labels (*fig. 14*, lines 14-19), and the format statement (*fig. 14*, line 20). Because the data printback option was not used (IECHO=0), the program prints only the first 20 data cards in raw and transformed versions (*fig. 14*, lines 21-55 and 57-65). This permits the user to check a sample to assure that the data are being transformed correctly. The observations summary (*fig. 14*, line 66) is followed by the proportional mortality observed in the controls (*fig. 14*, line 67).

The statistical portion of the printout begins with initial parameter estimates (*fig. 14*, line 68-71), the initial log likelihood values (*fig. 14*, line 72), iterations totals (*fig. 14*, lines 73-74) and the final log likelihood value (*fig. 14*, line 75). These precede the table of parameter values, their standard errors, and t-ratios (*fig. 14*, lines 76-81). In this example, the only parameter with a significant t-ratio is log(D); the values of the ratios for all other parameters fall below the critical $t = 1.96$ tabular value. The covariance matrix (*fig. 14*, lines 82-88) and prediction success table (*fig. 14*, lines 89-98) follow. The OPC indicates good

prediction success. Finally, the significance of the full model is tested (*fig. 14*, lines 99-100).

The restricted model [2] is computed next, with the $(\log W)^2$ parameter omitted (*fig. 14*, line 102). The initial estimates of the parameters that are retained are listed (*fig. 14*, lines 103-104), followed by the initial log likelihood value (*fig. 14*, line 105) and the iteration summary (*fig. 14*, lines 106-107). The final log likelihood value (*fig. 14*, line 108) and parameter estimates with their standard errors and t-ratios (*fig. 14*, lines 109-113) follow. In the restricted model, the t-ratios of all the parameters except natural response are now significant; this contrasts with the lack of significance of all parameters except dose in model [1]. The next portion of the printout is the usual presentation of the covariance matrix (*fig. 14*, lines 114-119), followed by the prediction success table (*fig. 14*, lines 120-129) and LR test of the hypothesis that the regression coefficients of the restricted model equal zero (*fig. 14*, lines 130-131). The hypothesis is rejected. Finally, the LR test of the hypothesis $H: \delta_3=0$, which was outlined in *section 4.3.1*, is presented (*fig. 14*, lines 132-133). The maximum log likelihood for the unrestricted model, the model including $(\log \text{weight})^2$, is $L(\Omega) = -93.6792$ and for the restricted model, the one excluding $(\log \text{weight})^2$, is $L(\omega) = -94.9086$. Since $LR = 2[L(\Omega) - L(\omega)] = 2(1.2294) = 2.4589 < 3.84$, the hypothesis H is accepted at the 0.05 significance level. Consequently, we conclude that $(\log \text{weight})^2$ is not a relevant explanatory variable in the regression.

Figure 14

```
1  WEIGHT STUDY,ROBERTSON ET AL,RANDOM DESIGN2,MEXACARBATE. HIGHER TERM Z IN MODEL POL02 28 AUG 80 09:32:55 4 1
2  CONTROL CARD: 4,4,0,2,1,0,0,0,0,3,0
3  NUMBER OF VARIABLES: 4
4  NUMBER OF VARIABLES ON INPUT CARDS: 4
5  NUMBER OF PARAMETERS: 5
6  TYPE OF MODEL: PROBIT
7  THERE ARE TRANSFORMATIONS
8  A CONTROL GROUP IS INDICATED BY VARIABLE 2 BEING ZERO
9  A RESTRICTED MODEL OF THE FIRST 3 VARIABLES IS TO BE COMPUTED ALSO
10 INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY
11 NATURAL RESPONSE IS A PARAMETER
12 ONE SUBJECT PER DATA CARD
13 TRANSFORMATIONS: 2L2=3L3=3L3L *4=
14 PARAMETER LABELS
15 1 CONSTANT
16 2 LOG(D)
17 3 LOG(W)
18 4 LOGWLOGW
19 5 NATURAL
```

Figure 14—Continued

20 FORMAT: (10X,F4.2,F5.1,T24,I1)

21 INPUT DATA

22	R2	1	1	001	000	679	2	0	
23	R2	1	1	002	000	861	2	0	
24	R2	1	1	003	000	762	2	0	
25	R2	1	1	004	000	635	2	0	
26	R2	1	1	005	000	973	2	0	
27	R2	1	1	006	000	604	2	0	
28	R2	1	1	007	000	963	2	0	
29	R2	1	1	008	000	745	2	0	
30	R2	1	1	009	005	760	2	0	
31	R2	1	1	010	005	791	2	0	
32	R2	1	1	011	005	766	2	0	
33	R2	1	1	012	005	930	2	0	
34	R2	1	1	013	005	727	2	0	
35	R2	1	1	014	005	614	2	0	
36	R2	1	1	015	005	838	2	0	
37	R2	1	1	016	010	815	2	0	
38	R2	1	1	017	010	728	2	0	
39	R2	1	1	018	010	764	2	0	
40	R2	1	1	019	010	599	2	1	
41	R2	1	1	020	010	652	2	1	
42	ETC.								

43 TRANSFORMED DATA

44	1.	0	1.00000	-1.30103	1.88081	3.53746
45	2.	0	1.00000	-1.30103	1.89818	3.60307
46	3.	0	1.00000	-1.30103	1.88423	3.55032
47	4.	0	1.00000	-1.30103	1.96848	3.87493
48	5.	0	1.00000	-1.30103	1.86153	3.46531
49	6.	0	1.00000	-1.30103	1.78817	3.19755
50	7.	0	1.00000	-1.30103	1.92324	3.69887
51	8.	0	1.00000	-1.00000	1.91116	3.65252
52	9.	0	1.00000	-1.00000	1.86213	3.46753
53	10.	0	1.00000	-1.00000	1.88593	3.55672
54	11.	1	1.00000	-1.00000	1.77743	3.15925
55	12.	1	1.00000	-1.00000	1.81425	3.29149

56 WEIGHT STUDY,ROBERTSON ET AL,RANDOM DESIGN2,MFXACARBATE. HIGHER TERM Z IN MODEL POLO2 28 AUG 80 09:32:57 6 2

57	13.	1	1.00000	-1.00000	1.84572	3.40668
58	14.	0	1.00000	-1.00000	1.80414	3.25492
59	15.	0	1.00000	-1.00000	1.83696	3.37441
60	16.	1	1.00000	-1.00000	1.77305	3.14372
61	17.	0	1.00000	-1.00000	1.81954	3.31074
62	18.	0	1.00000	-1.00000	1.82478	3.32981
63	19.	1	1.00000	-.698970	1.89763	3.60099
64	20.	0	1.00000	-.698970	1.91009	3.64845
65	ETC.					

66 207 OBSERVATIONS + 46 CONTROLS = 253 TOTAL
67 MORTALITY IN CONTROLS: .022

68 INITIAL ESTIMATES OF THE PARAMETERS

69	61.92830	CONSTANT	13.86360	LOGWLOGW
70	2.314442	LOG(D)	2.173913	-01 NATURAL
71	-58.12285	LOG(W)		

72 STARTING LIKELIHOOD = -97.5068

73 13 LOGIT EVALUATIONS

74 7 PROBIT EVALUATIONS

75 MAXIMUM LOG LIKELIHOOD = -93.6792

76	PARAMETER	STD ERROR	T RATIO
77	CONSTANT	99.4907	55.5623
78	LOG(D)	3.41625	.456533
79	LOG(W)	-94.6854	58.0475
80	LOGWLOGW	22.8966	15.1593
81	NATURAL	.144445-01	.142699-01

82 COVARIANCE MATRIX

83	CONSTANT	LOG(D)	LOG(W)	LOGWLOGW	NATURAL
84	3087.166	6.196166	-3224.037	841.1348	.1926991-01
85	6.196166	.2084226	-6.060385	1.509324	.5030425-03
86	-3224.037	-6.060385	3369.508	-879.6802	-.1834847-01
87	841.1348	1.509324	-879.6802	229.8032	.4319225-02
88	.1926991-01	.5030425-03	-.1834847-01	.4319225-02	.2036306-03

89 PREDICTION SUCCESS TABLE

90	PREDICTED OUTCOME	
91	-----	
92	ALIVE	DEAD
93	-----	

Figure 14—Continued

94	ALIVE	72	24
95	DEAD	23	88
96	-----		
97	PERCENT CORRECT	76	79
98	OVERALL PERCENT CORRECT: 77		

99 TESTING HYPOTHESIS THAT THE 4-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
 100 LIKELIHOOD RATIO TEST: CHI-SQUARED = 108.9986 DF=4 TAIL PROBABILITY = .000

101 WEIGHT STUDY, ROBERTSON ET AL, RANDOM DESIGN 2, MEXACARBALE, HIGHER TERM Z IN MODEL POL02 28 AUG 80 09:32:59 9 3

102 RESTRICTED MODEL WITH 3 VARIABLES AND 4 PARAMETERS

103 INITIAL ESTIMATES OF THE PARAMETERS

104 11.99335 CONSTANT 2.303810 LOG(D) -5.455093 LOG(W) .2173913-01 NATURAL

105 STARTING LIKELIHOOD = -98.2843

106 9 LOGIT EVALUATIONS

107 7 PROBIT EVALUATIONS

108 MAXIMUM LOG LIKELIHOOD = -94.9086

109	PARAMETER	STD ERROR	T RATIO
110	CONSTANT 15.9899	2.84374	5.62283
111	LOG(D) 3.30682	.430669	7.67834
112	LOG(W) -7.18592	1.43618	-5.00348
113	NATURAL .148626-01	.147205-01	1.00965

114 COVARIANCE MATRIX

115	CONSTANT	LOG(D)	LOG(W)	NATURAL
116	CONSTANT 8.086880	.5906353	-4.061704	.4634708-02
117	LOG(D) .5906353	.1854755	-.2444996	.4560381-03
118	LOG(W) -4.061704	-.2444996	2.062625	-.2457787-02
119	NATURAL .4634708-02	.4560381-03	-.2457787-02	.2166941-03

120 PREDICTION SUCCESS TABLE

121	ACTUAL OUTCOME	PREDICTED OUTCOME	
122		ALIVE	DEAD
123	ALIVE	72	24
124	DEAD	23	88
125	PERCENT CORRECT	76	79
126	OVERALL PERCENT CORRECT: 77		

130 TESTING HYPOTHESIS THAT THE 3-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
 131 LIKELIHOOD RATIO TEST: CHI-SQUARED = 106.5107 DF=3 TAIL PROBABILITY = .000

132 TESTING HYPOTHESIS THAT THE 4-BETA-PARAMETER MODEL AND THE 3-BETA-PARAMETER MODEL ARE THE SAME
 133 LIKELIHOOD RATIO TEST: CHI-SQUARED = 2.4589 DF=1 TAIL PROBABILITY = .117

4.4 Body Weight as a Variable: PROPORTIONAL Option

Dosage estimates are the primary objective of many toxicological investigations. The topical application technique described by Savin and others (1977), for example, is used to obtain precise estimates of the amounts of chemicals necessary to affect 50 or 90 percent of test subjects. The quality of chemical applied is known, as is the weight of test subjects. In the previous example, we tested the significance of a higher order term in a polynomial model. On the basis of an LR test, we conclude that the higher order term was not relevant. The PROPORTIONAL option permits the user to test the hypothesis that the response of test subjects is proportional to their body weight. If the hypothesis of proportionality is correct, LD₅₀ and LD₉₀ estimates at weights chosen by the investigator can be calculated.

4.4.1 Models and Hypotheses

We now consider the model

$$y = \delta_0 + \delta_1 x_1 + \delta_2 x_2$$

Where y = the probit or logit of the response, x_1 = logarithm of the dose (log D), and x_2 = logarithm of the weight (log W). Let $\delta_0 = \beta_0$, $\delta_1 = \beta_2$, and $\beta_2 = \delta_1 + \delta_2$. Then the model [1] can be rewritten as

$$y = \beta_0 + \beta_1 \log (D/ W) + \beta_2 \log W.$$

The hypothesis of proportionality is $H: \beta_2=0$ (Robertson and others 1981). The unrestricted model is [1] and the restricted model is

$$y = \beta_0 + \beta_1 \log (D/W)$$

The required maximum log likelihoods are $L(\Omega)$ and $L(\omega)$.

- (i) $L(\Omega)$: ML estimation of [2].
- (ii) $L(\omega)$: ML estimation of [3].

When the hypothesis H is true, asymptotically,

$$\text{LR} = 2[\text{L}(\Omega) - \text{L}(\omega)] \sim \chi^2 \quad (1)$$

so that the hypothesis H is accepted at the α significance level if

$$\text{LR} \leq \chi^2(1).$$

The necessary calculations are automatically performed when the PROPORTIONAL option is chosen. Note that the hypothesis can also be tested using the t-ratio for β_2 . If the hypothesis is accepted, the user may obtain LD₅₀ and LD₉₀ estimates for any body weight desired with the method described by Savin and others (1981). If the hypothesis is rejected, the BASIC option (*sec. 4.5*) should be used.

4.4.2 Input

The input (*fig. 15*) may begin with the usual starter cards (*fig. 15-A*) unless the PROPORTIONAL option is used after another analysis (except another PROPORTIONAL option set or the BASIC option—no other analysis may follow the use of either). In this example, input begins with the program cards (*fig. 15-B*).

The program card must have PROPORTIONAL in some position from columns 2-80 (card 6). The control card (card 7) must have NVAR=3 and KERNEL=2 so that the proportionality hypothesis will be tested. If the LR test of proportionality is not needed, NVAR=2 and KERNEL=0 may be specified; however, we suggest that the LR test be performed unless there is ample evidence that proportional response can be assumed. NPLL must be zero in either case, but other integers on the control card may vary as needed.

In this example, the transformations specified (card 8) are $y_2 = \log(x_2/x_3)$ and $y_3 = \log x_3$. If the LR test of proportionality is not requested and NVAR=2, KERNEL=0 are present on the control card, the transformation $y_2 = \log(x_2/x_3)$ alone- should be used. Parameters are: constant, log dose divided by log weight,

[illegible]

Figure 15

log weight, and natural response. Card 9 has labels suggestive of these names. The format card (card 10) instructs the program to read dose (fig. 15-C, columns 11-14), weight (fig. 15-C, columns 15-19), and response (fig. 15-C, column 24) from the data cards. The data cards are followed by an END card (card 264).

If the proportionality hypothesis is accepted, weights specified by the user may be placed behind the END card to obtain LD₅₀ and LD₉₀ estimates. One weight can be punched on each card in free field format.

4.4.3 Output

The output follows the usual pattern until the last portion. Titles appear at the top of each page (fig. 16, lines 1, 56 and 96). The descriptive section lists specification statements (fig. 16, line 3-12), transformation statement (fig. 16, line 13), parameter labels (fig. 16, lines 14-18), format statement (fig. 16, line 19), abbreviated raw and transformed data listing (fig. 16, lines 20-55 and 57-64), the observation summary (fig. 16, line 65), and the natural mortality statement (fig. 16, line 66).

The statistical portion of the printout begins, as usual, with the initial parameter estimates (fig. 16, lines 67-68), starting log likelihood value (fig. 16, line 69), and the iteration summary (fig. 16, lines 70-71) preceding the final log likelihood value (fig. 16, line 72). Parameter values and statistics (fig. 16, lines 73-77), covariance matrix (fig. 16, lines 78-83), prediction success table with OPC (fig. 16, lines 84-93) and test of significance of coefficients in the full model [2] follow (fig. 16, lines 94-95). The statistics and LR

test for the restricted model [3] are printed next (fig. 16, lines 97-124).

4.4.4 Hypothesis Testing

The LR statistic is

$$\begin{aligned} LR &= 2[L(\Omega) - L(\omega)] = 2[-94.9086 + 99.4884] \\ &= 2[4.5788] = 9.1596. \end{aligned}$$

The hypothesis of proportionality is rejected at the 0.05 significance level because the χ^2 critical value is 3.84. Note that the hypothesis is also rejected by the t-test because the t-ratio for β_2 is -2.92. The calculations in the last section of the printout are statistics based on average weight; these should be disregarded unless the proportionality hypothesis was accepted (see sec. 4.5.2 for an explanation of the printout).

4.5 Body Weight as a Variable: BASIC Option

The BASIC option estimates lethal doses D in the equation

$$y = \beta_0 + \beta_1(\log D) + \beta_2(\log W)$$

when $\beta_1 \neq \beta_2$. This model is appropriate when the proportionality hypothesis has been rejected, as in the previous example (sec. 4.4). Calculations are described by Savin and others (1981).

Figure 16

```

1  WEIGHT STUDY.R2 MEXACARBATE LOG(D/W),LOG(W) PROPORTIONAL          POL02    28 AUG 80  12:06:38    4    1

2  CONTROL CARD: 3,2,0,2,1,0,0,0,0,0,2,0

3  NUMBER OF VARIABLES: 3
4  NUMBER OF VARIABLES ON INPUT CARDS: 2
5  NUMBER OF PARAMETERS: 4
6  TYPE OF MODEL: PROBIT
7  THERE ARE TRANSFORMATIONS
8  A CONTROL GROUP IS INDICATED BY VARIABLE 2 BEING ZERO
9  A RESTRICTED MODEL OF THE FIRST 2 VARIABLES IS TO BE COMPUTED ALSO
10 INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY
11 NATURAL RESPONSE IS A PARAMETER
12 ONE SUBJECT PER DATA CARD

13 TRANSFORMATIONS: 23/L2=3L3=

14 PARAMETER LABELS
15   1  CONSTANT
16   2  LOG(D/W)
17   3  LOG(W)
18   4  NATURAL

19 FORMAT: (10X,F4.2,F5.1,T24,I1)

20 INPUT DATA
21 R2 1 1 001 000 674 2 0
22 R2 1 1 002 000 861 2 0
23 R2 1 1 003 000 762 2 0
24 R2 1 1 004 000 635 2 0
25 R2 1 1 005 000 973 2 0

```

Figure 16—Continued

```

26 R2 1 1 006 000 604 2 0
27 R2 1 1 007 000 963 2 0
28 R2 1 1 008 000 745 2 0
29 R2 1 1 009 005 760 2 0
30 R2 1 1 010 005 791 2 0
31 R2 1 1 011 005 766 2 0
32 R2 1 1 012 005 930 2 0
33 R2 1 1 013 005 727 2 0
34 R2 1 1 014 005 814 2 0
35 R2 1 1 015 005 838 2 0
36 R2 1 1 016 010 815 2 0
37 R2 1 1 017 010 728 2 0
38 R2 1 1 018 010 769 2 0
39 R2 1 1 019 010 599 2 1
40 R2 1 1 020 010 652 2 1
41 ETC.

```

42 TRANSFORMED DATA

```

43 1. 0 1.00000 -3.18184 1.88081
44 2. 0 1.00000 -3.19921 1.89818
45 3. 0 1.00000 -3.18526 1.88423
46 4. 0 1.00000 -3.26951 1.96848
47 5. 0 1.00000 -3.16256 1.86153
48 6. 0 1.00000 -3.08920 1.78817
49 7. 0 1.00000 -3.22427 1.92324
50 8. 0 1.00000 -2.91116 1.91116
51 9. 0 1.00000 -2.86213 1.86213
52 10. 0 1.00000 -2.88593 1.88593
53 11. 1 1.00000 -2.77743 1.77743
54 12. 1 1.00000 -2.81425 1.81425
55 13. 1 1.00000 -2.84572 1.84572

```

56 WEIGHT STUDY, R2 MEXACARBATE LOG(D/W), LOG(W) PROPORTIONAL

POL02 28 AUG 80 12:06:42 6 2

```

57 14. 0 1.00000 -2.80414 1.80414
58 15. 0 1.00000 -2.83696 1.83696
59 16. 1 1.00000 -2.77305 1.77305
60 17. 0 1.00000 -2.81954 1.81954
61 18. 0 1.00000 -2.82478 1.82478
62 19. 1 1.00000 -2.59660 1.89763
63 20. 0 1.00000 -2.60906 1.91009
64 ETC.

```

65 207 OBSERVATIONS + 46 CONTROLS = 253 TOTAL
66 MORTALITY IN CONTROLS: .022

67 INITIAL ESTIMATES OF THE PARAMETERS
68 12.22245 CONSTANT

2.269062 LOG(D/W) -3.317775 LOG(W) .2173913-01 NATURAL

69 STARTING LIKELIHOOD = -98.4603

70 12 LOGIT EVALUATIONS

71 7 PROBIT EVALUATIONS

72 MAXIMUM LOG LIKELIHOOD = -94.9086

73	PARAMETER	STD ERROR	T RATIO
74	CONSTANT 15.9899	2.84379	5.62273
75	LOG(D/W) 3.30682	.430670	7.67832
76	LOG(W) -3.87910	1.32634	-2.92467
77	NATURAL .148627-01	.147207-01	1.00965

78 COVARIANCE MATRIX

79	CONSTANT	LOG(D/W)	LOG(W)	NATURAL
80	CONSTANT 8.087159	.5906401	-3.471210	.4634957-02
81	LOG(D/W) .5906401	.1854768	-.5902480-01	.4560508-03
82	LOG(W) -3.471210	-.5902480-01	1.759175	-.2001868-02
83	NATURAL .4634957-02	.4560508-03	-.2001868-02	.2166987-03

84 PREDICTION SUCCESS TABLE

85		PREDICTED OUTCOME	
86		-----	
87	ACTUAL OUTCOME	ALIVE	DEAD
88		-----	-----
89	ALIVE	72	24
90	DEAD	23	88
91		-----	-----
92	PERCENT CORRECT	76	79
93	OVERALL PERCENT CORRECT: 77		

94 TESTING HYPOTHESIS THAT THE 4-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
95 LIKELIHOOD RATIO TEST: CHI-SQUARED = 106.5108 DF=3 TAIL PROBABILITY = .000

Figure 16—Continued

```

96  WEIGHT STUDY.R2 MEXACARBATE LOG(D/W),LOG(W) PROPORTIONAL      POL02      28 AUG 80  12:06:44   8   3

97  RESTRICTED MODEL WITH 2 VARIABLES AND 3 PARAMETERS

98  INITIAL ESTIMATES OF THE PARAMETERS
99      7.334388 CONSTANT          2.791762 LOG(D/W)          .2173913-01 NATURAL

100  STARTING LIKELIHOOD = -100.4000

101  10 LOGIT EVALUATIONS
102  11 PRORIT EVALUATIONS

103  MAXIMUM LOG LIKELIHOOD = -99.4884

104  PARAMETER      STD ERROR      T RATIO
105  CONSTANT      8.59942      1.06703      8.05919
106  LOG(D/W)      3.28190      .412611     7.95398
107  NATURAL      .139061-01     .137678-01     1.01004

108  COVARIANCE MATRIX
109      CONSTANT      LOG(D/W)      NATURAL
110  CONSTANT      1.138559      .4380162     .5864396-03
111  LOG(D/W)      .4380162      .1702479     .3247294-03
112  NATURAL      .5864396-03     .3247294-03     .1895527-03

113  PREDICTION SUCCESS TABLE
114      PREDICTED OUTCOME
115      -----
116  ACTUAL OUTCOME      ALIVE      DEAD
117  -----
118  ALIVE                68        28
119  DEAD                 19        92
120  -----
121  PERCENT CORRECT          78        77
122  OVERALL PERCENT CORRECT: 77

123  TESTING HYPOTHESIS THAT THE 2-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
124  LIKELIHOOD RATIO TEST: CHI-SQUARED = 97.4199  DF=2  TAIL PROBABILITY = .000

125  PROPORTIONAL MODEL

126  WBAR=77.367111, LOG10(WBAR)=1.888556

127  A=2.4013643, B=3.2819001

128  V11=.0913368, V12=.1164934, V22=.1702479

129  SE(A)=.3022199, SE(B)=.4126111

130  G=.060714, T=1.95996

131  LD50=.1854815, LIMITS=.1576298,.2151935

132  LD90=.4558178, LIMITS=.3707691,.6206283

```

4.5.1 Input

The input (fig. 17) begins with the usual starter cards (fig. 17-A) unless the BASIC option follows another analysis (except another BASIC or PROPORTIONAL set). In that case, input would begin with the program cards (fig. 17-B). The program cards must have:

1. BASIC is some position on the title card (card 6).
2. A control card (card 7) with NVAR=3, NPLL=0, and KERNEL=0. The basic model is limited to three regression coefficients (NVAR), no parallel groups, and no test of a restricted model. The other integers may vary, as required.

In this example, the transformation card (card 8) states that $y_2 = \log x_2$ and $y_3 = \log x_3$. The parameters (card 9) are labeled: "constant," "log(D)," and "log(W)." Natural response is a parameter in this example, but need not be present in each use of the BASIC option. The format statement (card 10) instructs the program to read only dose (fig. 17-C, columns 11-14), body weight (fig. 17-C, columns 15-19), and response (fig. 17-C, column 24). The data are followed by an END card (fig. 17-D, card 264) and weight cards (fig. 17-D, cards 265-269).

4.5.2 Output

The BASIC output follows the usual pattern until the last section. Title repetition on each page (fig. 18, lines 1, 56, 99, and 125), control card listing (fig. 18, line 2), specification statements (fig. 18, lines 3-11), transformation listing (fig. 18, line 12), parameter labels (fig. 18,

lines 13-17), format statement listing (fig. 16, line 18), an abbreviated raw and transformed data listing (fig. 18, lines 19-55 and 57-63), observations summary (fig. 18, line 64), and natural response statement (fig. 18, line 65) form the descriptive portion of the output.

The statistical portion contains the usual initial parameter estimates (fig. 18, lines 66-67), initial log likelihood value (fig. 18, line 68), iteration summary (fig. 18, lines 69-70) final log likelihood value (fig. 18, line 71), parameter values and statistics (fig. 18, lines 72-76), covariance matrix (fig. 18, lines 77-82), prediction success table (fig. 18, lines 83-92), and LR test of the full model (fig. 18, lines 93-94).

Statistics for the model using average weight of the test subjects as the value of W follows (fig. 18, lines 95-98 and 100-102). The terminology of these statistics is as follows. WBAR is average weight, and LOGIO (WBAR) is the logarithm of WBAR to the base 10 (fig. 18, line 96). The parameters are called "A" and "B"; A is the intercept of the line calculated at WBAR, and B is the slope of the line (fig. 18, line 97). The variances and covariances of the parameters are listed next (fig. 18, line 98), followed by the standard errors of intercept and slope (fig. 18, line 100). Values of g and t , used to calculate confidence limits (Finney 1971) for point estimates on a probit or logit line, appear next (fig. 16, line 101). Finally, values of the lethal dose necessary for 50 and 90 percent mortality at WBAR, together with their 95 percent confidence limits, are printed (fig. 18, lines 101-102). Next, statistics for each weight specified on the weight cards (fig. 17-D, cards 265-269) are printed (fig. 18, lines 104-124 and 126-139).

1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30	31	32	33	34	35	36	37	38	39	40	41	42	43	44	45	46	47	48
@RUN,8IRV,143-8801,,3,100POLO2																																															
@ASG,TX.,TU20009																																															
@COPY,GX.																																															
@FREEX.																																															
@XQTPOLO2																																															
=WEIGHT STUDY.R2 MEXACARBATE LOG(D)LOG(W)BASIC																																															
3,2,0,2,1,0,0,0,0,0,0,0,0																																															
2L2=3L3=																																															
CONSTANTLOG(D)LOG(W)NATURAL																																															
(10X,F4.2,F5.1,T24,I1)																																															
R21100100067920																																															

263	R216098050766241																																															
264	END																																															
265	60																																															
266	70																																															
267	80																																															
268	90																																															
269	100																																															

Figure 17.

Figure 18

1	WEIGHT STUDY.R2 MEXACARBATE LOG(D) LOG(W) BASIC	POLO2	17 SEP 80	12:55:39	4	1
2	CONTROL CARD: 3,2,0,2,1,0,0,0,0,0,0					
3	NUMBER OF VARIABLES: 3					
4	NUMBER OF VARIABLES ON INPUT CARDS: 2					
5	NUMBER OF PARAMETERS: 4					
6	TYPE OF MODEL: PROBIT					
7	THERE ARE TRANSFORMATIONS					
8	A CONTROL GROUP IS INDICATED BY VARIABLE 2 BEING ZERO					
9	INITIAL ESTIMATES OF THE PARAMETERS ARE TO BE CALCULATED AUTOMATICALLY					
10	NATURAL RESPONSE IS A PARAMETER					
11	ONE SUBJECT PER DATA CARD					
12	TRANSFORMATIONS: 2L2=3L3=					
13	PARAMETER LABELS					
14	1 CONSTANT					
15	2 LOG(D)					
16	3 LOG(W)					
17	4 NATURAL					
18	FORMAT: (10X,F4.2,F5.1,T24,I1)					
19	INPUT DATA					
20	R2 1 1 001 000 679 2 0					
21	R2 1 1 002 000 861 2 0					
22	R2 1 1 003 000 762 2 0					
23	R2 1 1 004 000 635 2 0					
24	R2 1 1 005 000 973 2 0					
25	R2 1 1 006 000 604 2 0					
26	R2 1 1 007 000 963 2 0					
27	R2 1 1 008 000 745 2 0					
28	R2 1 1 009 005 760 2 0					
29	R2 1 1 010 005 791 2 0					
30	R2 1 1 011 005 766 2 0					
31	R2 1 1 012 005 930 2 0					
32	R2 1 1 013 005 727 2 0					
33	R2 1 1 014 005 614 2 0					
34	R2 1 1 015 005 838 2 0					
35	R2 1 1 016 010 815 2 0					
36	R2 1 1 017 010 728 2 0					
37	R2 1 1 018 010 769 2 0					
38	R2 1 1 019 010 599 2 1					
39	R2 1 1 020 010 652 2 1					
40	ETC.					
41	TRANSFORMED DATA					
42	1. 0 1.00000 -1.30103 1.88081					
43	2. 0 1.00000 -1.30103 1.89818					
44	3. 0 1.00000 -1.30103 1.88423					
45	4. 0 1.00000 -1.30103 1.96848					
46	5. 0 1.00000 -1.30103 1.86153					
47	6. 0 1.00000 -1.30103 1.78817					
48	7. 0 1.00000 -1.30103 1.92324					
49	8. 0 1.00000 -1.00000 1.91116					
50	9. 0 1.00000 -1.00000 1.86213					
51	10. 0 1.00000 -1.00000 1.88593					
52	11. 1 1.00000 -1.00000 1.77743					
53	12. 1 1.00000 -1.00000 1.81425					
54	13. 1 1.00000 -1.00000 1.84572					
55	14. 0 1.00000 -1.00000 1.80414					
56	WEIGHT STUDY.R2 MEXACARBATE LOG(D) LOG(W) BASIC	POLO2	17 SEP 80	12:55:41	6	2
57	15. 0 1.00000 -1.00000 1.83696					
58	16. 1 1.00000 -1.00000 1.77305					
59	17. 0 1.00000 -1.00000 1.81954					
60	18. 0 1.00000 -1.00000 1.82478					
61	19. 1 1.00000 -.698970 1.89763					
62	20. 0 1.00000 -.698970 1.91009					
63	ETC.					
64	207 OBSERVATIONS + 46 CONTROLS = 253 TOTAL					
65	MORTALITY IN CONTROLS: .022					
66	INITIAL ESTIMATES OF THE PARAMETERS					
67	11.99335 CONSTANT 2.303810 LOG(D) -5.455093 LOG(W) .2173913-01 NATURAL					
68	STARTING LIKELIHOOD = -98.2843					
69	9 LOGIT EVALUATIONS					
70	7 PROBIT EVALUATIONS					

Figure 18—Continued

71 MAXIMUM LOG LIKELIHOOD = -94.9086

	PARAMETER	STD ERROR	T RATIO
72	CONSTANT	15.9899	2.84374
73	LOG(D)	3.30682	.430669
74	LOG(W)	-7.18592	1.43618
75	NATURAL	.148626-01	.147205-01

77 COVARIANCE MATRIX

	CONSTANT	LOG(D)	LOG(W)	NATURAL
78	CONSTANT	.5906353	-4.061704	.4634708-02
79	LOG(D)	.1854755	-.2444996	.4560381-03
80	LOG(W)	-4.061704	-.2444996	2.062625
81	NATURAL	.4634708-02	.4560381-03	-.2457787-02
82				.2166941-03

83 PREDICTION SUCCESS TABLE

ACTUAL OUTCOME	PREDICTED OUTCOME	
	ALIVE	DEAD
84		
85		
86		
87		
88	72	24
89	23	88
90		
91	76	79
92	OVERALL PERCENT CORRECT: 77	

93 TESTING HYPOTHESIS THAT THE 3-BETA-PARAMETER MODEL AND THE 0-BETA-PARAMETER MODEL ARE THE SAME
94 LIKELIHOOD RATIO TEST: CHI-SQUARED = 106.5107 DF=3 TAIL PROBABILITY = .000

95 BASIC MODEL

96 WBAR=77.367111, LOG10(WBAR)=1.888556

97 A=2.4188659, B=3.3068204

98 V11=.1020167, V12=.1288841, V22=.1854755

99 WEIGHT STUDY.R2 MEXACARBATE LOG(D) LOG(W) BASIC

POL02

17 SEP 80 12:55:44

8

3

100 SE(A)=.3194006, SE(B)=.4306687

101 G=.065157, T=1.95996

102 LD50=.1855761, LIMITS=.1573355,.2163299

103 LD90=.4529707, LIMITS=.3654136,.6284598

104 W=60.000000, LOG10(W)=1.778151

105 A=3.212284, B=3.3068204

106 V11=.1638840, V12=.1558781, V22=.1854755

107 SE(A)=.4048259, SE(B)=.4306687

108 G=.065157, T=1.95996

109 LD50=.1068083, LIMITS=.0799298,.1368499

110 LD90=.2607072, LIMITS=.2007445,.3676450

111 W=70.000000, LOG10(W)=1.845098

112 A=2.7311541, B=3.3068204

113 V11=.1203684, V12=.1395096, V22=.1854755

114 SE(A)=.3469415, SE(B)=.4306687

115 G=.065157, T=1.95996

116 LD50=.1493085, LIMITS=.1230761,.1768952

117 LD90=.3644454, LIMITS=.2943531,.4990451

Figure 18—Continued

118	W=80.000000, LOG10(W)=1.903090				
119	A=2.5144286, B=3.3068204				
120	V11=.0976179, V12=.1253306, V22=.1854755				
121	SE(A)=.3124386, SE(B)=.4306687				
122	G=.065157, T=1.95996				
123	LD50=.1995743, LIMITS=.1692023,.2335790				
124	LD90=.4871386, LIMITS=.3905918,.6827091				
125	WEIGHT STUDY.R2 MEXACARBATE LOG(D) LOG(W) BASIC	POL02	17 SEP 80	12:55:44	9 4
126	W=90.000000, LOG10(W)=1.954243				
127	A=1.9468505, B=3.3068204				
128	V11=.0890662, V12=.1128239, V22=.1854755				
129	SE(A)=.2984396, SE(B)=.4306687				
130	G=.065157, T=1.95996				
131	LD50=.2577876, LIMITS=.2112662,.3165341				
132	LD90=.6292309, LIMITS=.4848818,.9305364				
133	W=100.000000, LOG10(W)=2.000000				
134	A=1.6180408, B=3.3068204				
135	V11=.0905631, V12=.1016362, V22=.1854755				
136	SE(A)=.3009370, SE(B)=.4306687				
137	G=.065157, T=1.95996				
138	LD50=.3241126, LIMITS=.2507269,.4269428				
139	LD90=.7911227, LIMITS=.5783192,1.2488830				
140	oFIN				

5. ERROR MESSAGES

Error messages clearly indicate mistakes in the input. For example:

Message	Reason
CONTROL CARD: 4,4,0,2,1,0,0,0,0,3,0 THERE SHOULD BE 12 NUMBERS, NOT 11 FORMAT: (10x,F4.2,F5.I(T24,I1))	One of the integers is missing from the control card. The parenthesis preceding T24 should have been a comma.

ILLEGAL FORMAT
CHARACTERS WERE
ACCEPTED AS BLANKS

TRANSFORMATIONS:
L2=3L3=3L3L3*4=
STACK MUST BE EMPTY AT
END OF POLISH STRING

CONTROL CARD:
4,4,0,2,1,0,0,0,0,3,0,0
THERE ARE NOW
16750372454
PARALLEL GROUPS,
EXCEEDING 3

An extra "3" was punched ("33" should be "3"), so an extra x_3 was put into the stack. The stack, therefore, was not empty at the end of the scan. See *section 2.4.1*.

The 3 belonged in the KERNEL column, not in the NPLL column.

Along with the error message, the user will receive a message guaranteed to catch the eye.

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Russell, Robert M.; Savin, N.E.; Robertson, Jacqueline L. **POLO2: a user's guide to multiple Probit Or LOGit analysis.** Gen. Tech. Rep. PSW-55. Berkeley, CA: Pacific Southwest Forest and Range Experiment Station, Forest Service, U.S. Department of Agriculture; 1981. 37 p.

This guide provides instructions for the use of POLO2, a computer program for multivariate probit or logic analysis of quantal response data. As many as 3000 test subjects may be included in a single analysis. Including the constant term, up to nine explanatory variables may be used. Examples illustrating input, output, and uses of the program's special features for hypothesis testing are included.

Retrieval terms: multiple probit analysis, multiple logit analysis, multivariate analysis