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MIX: A Computer Program to Evaluate Interaction Between Chemicals

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A computer program, MIX, was designed to identify pairs of chemicals whose interaction results in a response that departs significantly from the model predicated on the assumption of independent, uncorrelated joint action. This report describes the MIX program, its statistical basis, and instructions for its use.

Retrieval Terms: mixtures, synergism, antagonism

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INTRODUCTION

In the absence of biological data, elaborate theories and statistical procedures have been developed to explain the interaction of chemicals applied in a mixture to a target organism. Since Bliss (1939) first described statistical tests of joint action, complex methods have been devised to describe various hypothetical types of chemical interaction (e.g., Plackett and Hewlett 1948, Hewlett and Plackett 1950). These methods have not been widely used in entomological investigations because they are mathematically complex, difficult to interpret, and generally irrelevant to the practical consideration of identifying mixtures that are significantly more effective together than one would expect from their individual toxicities.

For general purposes, only three types of joint action between chemicals need be defined: synergism, antagonism and independence. We define synergism as activity significantly greater than that expected from the toxicities of the separate mixture components. Antagonism is activity that is significantly less than expected. Independence occurs when activity of the mixture does not differ significantly from that predicted on the basis of effectiveness of the separate components.

We designed a computer program called MIX to identify pairs of chemicals whose interaction results in a response that departs significantly from those predicted by the model of independent, uncorrelated joint action (Bliss 1939). In this model, toxicity of one component is assumed to be unaffected by the presence of the other chemical, and susceptibility to one chemical is not correlated with susceptibility to the other. Rejection of the model may indicate some correlation of susceptibilities, non-independent interaction, or both. The statistical procedure used in MIX is flexible: either a response range or a single response level may be evaluated.

This report describes the MIX computer program, its statistical basis, and instructions for its use.

STATISTICAL MODEL

The statistical model used in MIX was described by Robertson and Smith (1984) as follows:

Under the null hypothesis of independent, uncorrelated action, the probability $(1 - P_0)$ of surviving an exposure to a chemical mixture comprised of concentration X_1 of chemical 1 and X_2 of chemical 2 is the product of the probability of survival from natural causes $(1 - m)$, the probability of surviving $X_1(1 - P_1)$ and the probability of surviving $X_2(1 - P_2)$. Hence,

$$1 - P_0 = (1 - m)(1 - P_1)(1 - P_2)$$

or

$$P_0 = m + \{(1 - m)[P_1 + (1 - P_1)P_2]\}$$

If the true probability of response to the mixture is denoted by P , then the null hypothesis can be expressed as $H_0: P = P_0$.

For an actual experiment with a mixture of two chemicals, a dose X of the mixture is made up of X_1 of chemicals 1 and X_2 of chemical 2 in accordance with the mixing proportions used. Of n insects exposed to the mixture at concentration X , an observed proportion, $\hat{P}$, is killed. The expected response rate under the null hypothesis, P_0 , is estimated as

$$\hat{P}_0 = \hat{m} + (1 - \hat{m})[\hat{P}_1 + (1 - \hat{P}_1)\hat{P}_2]$$

where m is the proportional mortality of the n_0 subjects in the control group, and $\hat{P}_1$ and $\hat{P}_2$ are the predicted response probabilities to X_1 and X_2 , adjusted for natural mortality and estimated from the probit regression lines for chemicals 1 and 2 separately.

The difference $(\hat{P} - \hat{P}_0)$ is standardized by an estimate of the inherent variability associated with P and P_0 to assure that the magnitude of the difference is not the result of chance alone. Because $\hat{P}$ and $\hat{P}_0$ are statistically independent,

$$\begin{aligned}\text{Var}(\hat{P} - \hat{P}_0) &= \text{Var}(\hat{P}) + \text{Var}(\hat{P}_0) \\ &= P(1 - P)/n \\ &\quad + (1 - P_0)^2[1 + C_m)(1 + C_1)(1 + C_2) - 1]\end{aligned}$$

where $C_m = \text{Var}(\hat{m})/(1 - m)^2$, $C_1 = \text{Var}(\hat{P}_1)/(1 - \hat{P}_1)^2$ and

$C_2 = \text{Var}(\hat{P}_2)/(1 - P_2)^2$. Under the null hypothesis, this variance can be estimated by

$$\begin{aligned}\text{estimated Var}(\hat{P} - \hat{P}_0) &= \hat{P}_0(1 - \hat{P}_0)/n \\ &\quad + (1 - \hat{P}_0)^2[(1 + \hat{C}_m)(1 + \hat{C}_1)(1 + \hat{C}_2) - 1]\end{aligned}$$

Calculations of $\hat{C}_i$ require the estimated intercept $\hat{\beta}_{1i}$, slope $\hat{\beta}_{2i}$, and their estimated variances and covariance, $\hat{\sigma}_{1i}$, $\hat{\sigma}_{2i}$, $\hat{\sigma}_{12i}$ from the prohibit analysis of chemical i alone, $i = 1, 2$. Then $\hat{Y}_i = \hat{\beta}_{1i} + \hat{\beta}_{2i} \log_{10}(X_i)$, $V_i = \hat{\sigma}_{1i} + 2\hat{\sigma}_{12i} \log_{10} X_i + \hat{\sigma}_{2i} \log_{10}^2(X_i)$ and $\hat{C}_i = \exp(-Y_i^2/V_i) / [2\pi(1 - \hat{P}_i)^2]$; $\hat{C}_m = \hat{m}/[n_0(1 - \hat{m})]$.

The test statistic at dose X is $Z^2(X) = \frac{(P - P_0)^2}{\text{estimated Var}(\hat{P} - \hat{P}_0)}$

Under H_0 , if the number of observations in the control group at dose X of the mixture and at all doses of the individual components are all large, the central limit theorem assures that $Z^2(X)$ will be well approximated by a χ^2 distribution with one degree of freedom [$\chi^2_{(1)}$]. As long as all doses are located to avoid very high (>90 pct) and very low (<10 pct) response probabilities, sample sizes of 30 per dose should suffice for purposes of the approximation. The null hypothesis is rejected by an α -level test if $Z^2(X)$ exceeds the $(1 - \alpha)$ 100 percent quantile of the $\chi^2_{(1)}$ distribution.

As an approximate procedure for evaluating the toxicity of a mixture over a range of doses, the $Z^2(X)$ are summed over the k concentrations lying in the range of interest,

$$Z^2 = \sum_{j=1}^k Z^2(X_j)$$

The hypothesis of independent, uncorrelated action by the mixture over the entire range is rejected if Z^2 exceeds the $(1-\alpha)$ 100 percent quantile of the χ^2 distribution with k degrees of freedom.

DATA INPUT FILE

Parameter estimates from statistical analyses with each component and the mixture are placed in a data file for use by MIX. The file name is arbitrary, but must include 10 or fewer characters and no blanks. The name of the file is MIXIN (*fig. 1*).

Line 1 of the data file identifies the first chemical in the mixture (the choice is arbitrary, since either chemical may be considered number 1). Line 17 (*fig. 1*) is another example for a second data set. This identifier, a maximum of 10 characters, must be typed within single quotation marks (e.g., 'permethrin').

Line 2 (*fig. 1*) lists seven parameter estimates or data related to the insects' response to chemical 1. Line 18 is another example for a second data set. The values must be separated by

one or more blank spaces and must be entered in specific order. Parameter estimates are available from the printout of a probit or logit analysis program such as POLO (Russell and others 1977). In *figure 2*, the parameters are labelled as:

BI: estimated intercept
 B2: estimated slope
 S1: estimated variance of B 1
 S2: estimated variance of B2
 S12: estimated covariance of BI and B2
 ADJ: heterogeneity adjustment factor
 PROP: Proportion (as a decimal) of chemical 1

The next line (lines 3 and 19) identifies the second chemical in the mixture, entered as described for the first chemical. In the next line (lines 4, 20), the seven parameter estimates or data for the second chemical in the mixture are entered in the same format as values for the first chemical.

Data for the actual bioassay done with the mixture of chemical 1 and chemical 2 follow. The three entries (K, EXP, DEAD) in the first data line (lines 5, 21) show the number of doses or concentration levels tested (K) not counting the control group(s), number tested in the control group (EXP), and the number dead

```

1.      'permethrin'
2.      2.53 5.09 .1097 .6959 .2483 1.1825 .1
3.      'dimilin'
4.      1.43 1.47 .015 .0392 .0156 2.2685 .9
5.      11 69 13
6.      .001 10 6
7.      .01 10 5
8.      .1 68 29
9.      .2 60 38
10.     .3 60 48
11.     .5 59 48
12.     .7 60 53
13.     1. 70 69
14.     2. 60 60
15.     3. 60 60
16.     10. 10 10
17.     'permethrin'
18.     2.66 4.77 .0695 .2580 .1219 1.4107 .1
19.     'methoprene'
20.     1.4 1.64 .0134 .0367 .0151 1.6960 .9
21.     10 68 10
22.     .001 10 0
23.     .01 68 17
24.     .02 60 19
25.     .03 60 22
26.     .05 60 26
27.     .07 57 38
28.     .1 70 51
29.     .5 60 56
30.     1. 70 64
31.     10. 10 10
  
```

Figure 1--Contents of a sample input file named 'MIXIN.'

in the control group (DEAD). Following this line, data for each dose or concentration of the mixture are entered. Values listed in each line (6-10 and 21-31, *fig. 1*) are the dose (concentration) in actual units tested (DOSE), number of subjects exposed to the dose (EXP), and number of deaths observed among the subjects tested (DEAD). The format is the same for all of these entries; the three entries on each line must be in the order listed above and values must be separated by one or more blank spaces. The data file needs no special character to identify the end of the file.

RUNNING THE MIX PROGRAM

The MIX program (see *Appendix*) is written in Fortran 77. It consists of a main routine (MIXTURES. F77) and a subroutine (PINORM. 77). PINORM. 77 computes (X) for any X. All computations in both routines are done in double precision. The executable routine is MIXTURES.PR this program should be copied or moved into the directory that contains your input file.

After entering the directory where MIXTURES.PR and the input file are located, type X MIXTURES and hit the return key.

This message will appear:

ENTER FILE NAME WHERE DATA IS STORED.

The name of the input file should be typed at the end of this line; the name must be placed within single quotes. Hit the return key. For example, the completed line for the sample data file MIXIN would appear as

ENTER FILE NAME WHERE DATA IS STORED. 'MIXIN'

Output from MIX will be stored in the working directory. This file is created by the program if it does not already exist. On the next line, you will be asked to supply the name of the output file. The prompt line that appears is

ENTER FILE NAME TO STORE OUTPUT.

The name chosen should contain no more than 10 characters and no blanks (e.g., 'MIXOUT'). The name should be typed and

DIET TEST. PERMETHRIN + METHOPRENE

INTERCEPTS AND SLOPES UNCONSTRAINED. PREPARATION IS (1) PERMETHRIN
ESTIMATING NATURAL RESPONSE

MAXIMUM LOG-LIKELIHOOD -143.30810

	PARAMETER	STANDARD ERROR	T RATIO
PERMETHR	2.6591623	.26359936	10.087894
NATURAL	.58769126-01	.20471423-01	2.8707878
SLOPE	4.7721107	.50788916	9.3959688

VARIANCE-COVARIANCE MATRIX

	S1 PERMETHR	NATURAL	SLOPE	
PERMETHR	.6948462-01	.6208465-03	.1219397	← S12
NATURAL	.6208465-03	.4190794-03	.2681468-02	
SLOPE	.1219397	.2681468-02	.257914	← S2

CHI-SQUARED GOODNESS OF FIT TEST

PREPARATION	SUBJECTS	RESPONSES	EXPECTED	DEVIATION	PROBABILITY
PERMETHR	10	1.	.588	.412	.058769

CHI-SQUARE	38.089	DEGREES OF FREEDOM	27	HETEROGENEITY	1.4107 ← ADJ
------------	--------	--------------------	----	---------------	--------------

Figure 2--POLO output from analysis of bioassay data with individual chemicals.

1.	chemical		permethrin	dimilin	
2.	proportion		.1000	.9000	
3.	intercept b1		2.5300	1.4300	
4.	slope b2		5.0900	1.4700	
5.	var(b1)		.1097	.0150	
6.	var(b2)		.6959	.0392	
7.	covar(b1,b2)		.2483	.0156	
8.	het adjust		1.1825	2.2685	
9.	mixture of permethrin and dimilin				
number exposed	dose	no_exp	obs.mort	exp.mort	chisq
10.	controls	69.	.1884		
11.					
12.	.0010	10.	.6000	.1893	6.4324
13.	.0100	10.	.5000	.2350	2.4012
14.	.1000	68.	.4265	.5595	1.7354
15.	.2000	60.	.6333	.7008	.6445
16.	.3000	60.	.8000	.7758	.1299
17.	.5000	59.	.8136	.8549	.4823
18.	.7000	60.	.8833	.8960	.0661
19.	1.0000	70.	.9857	.9302	4.2753
observed	2.0000	60.	1.0000	.9756	3.3673
proportional	3.0000	60.	1.0000	.9913	1.9223
mortality	10.0000	10.	1.0000	1.0000	.2109
22.					
23.	chisq= 21.6676 with 11 degrees of freedom				
			sum of chi square contributions		
24.	chemical		permethrin	methoprene	
25.	proportion		.1000	.9000	
26.	intercept b1		2.6600	1.4000	
27.	slope b2		4.7700	1.6400	
28.	var(b1)		.0695	.0134	
29.	var(b2)		.2580	.0367	
30.	covar(b1,b2)		.1219	.0151	
31.	het adjust		1.4107	1.6960	
32.	mixture of permethrin and methoprene				
33.	dose	no_exp	obs.mort	exp.mort	chisq
34.	controls	68.	.1471		
35.	.0010	10.	.0000	.1472	11.7492
36.	.0100	68.	.2500	.1686	1.3341
37.	.0200	60.	.3167	.2084	1.7025
38.	.0300	60.	.3667	.2498	1.6851
39.	.0500	60.	.4333	.3256	1.2521
40.	.0700	57.	.6667	.3898	8.4009
41.	.1000	70.	.7286	.4681	9.1009
42.	.5000	60.	.9333	.8270	6.1600
43.	1.0000	70.	.9143	.9224	.0429
44.	10.0000	10.	1.0000	1.0000	.2939
45.	chisq= 41.7218 with 10 degrees of freedom				

Statistics from POLO output, for each chemical

Exected proportion killed assuming the hypothesis of independent uncorrelated joint action

contribution to chi square

Figure 3--Copy of output stored in the file named 'MIXOUT.'

entered, then hit the return key. For this example, the complete line would be

ENTER FILE NAME TO STORE OUTPUT. 'MIXOUT'

Upon receipt of these file names, MIX will process the data in the input files, then place the results into the output file. When the process is complete, STOP will appear on the screen.

To view the contents of the output file, enter TYPE followed by at least one space and the name of the output file. For the example shown in this manual, the complete line would be

TYPE MIXOUT

Hard copy of this output is obtained by typing QPRINT followed by at least one space and the name of the output file.

SAMPLE OUTPUT

Sample input data listed in *figure 1*, were subjected to probit analysis (*fig. 2*). Relevant output was stored in a file called MIXOUT (*fig. 3*). Data and statistical parameters entered for the individual chemicals are listed in lines 1-8 for permethrin and dimilin and lines 24-31 for permethrin and methoprene.

Next, MIX calculations of expected mortality and χ^2 values for each dose calculated as described in section 1 (Statistical Model) for the null hypothesis of independent joint action are listed. In this sample output, these calculations appear in the last two columns of lines 10-22 and 34-44. The total χ^2 with the appropriate degrees of freedom (number of doses [including the control]-1) are listed next (line 23, line 45). To determine whether significant departures from the null hypothesis level of choice (e.g., $P=0.05$) can be located in a χ^2 table available in any statistical text. If the value calculated by MIX is less than the tabular value, the null hypothesis cannot be rejected. If the χ^2 value is greater than the tabular value, then the hypothesis of independent joint action is rejected.

A plot of the data should be prepared to determine the direction of significant departures of the results from values predicted by the hypothesis. Ideally, all the data points will lie to the left (significant antagonism) or the right (significant synergism) of expected values. However, points on either side of the expected line occur (Robertson and Smith 1984): in these instances, χ^2 values above and below 50 percent may be examined to assist in interpretation of the results. Degrees of freedom (lines 23, 45) are equal to the number of doses (including controls) minus 1; if a subset of data is considered, degrees of freedom are calculated in the same way. The calculated χ^2 value (lines 23, 45) should be compared with tabular value at the appropriate degrees of freedom and probability value of choice to determine significance of the results.

APPENDIX: PROGRAM LISTING

c MIXTURES.F77 CONDUCTS A CHISQUARE GOODNESS OF FIT TEST OF THE
c HYPOTHESIS OF INDEPENDENT, UNCORRELATED JOINT ACTION BY COMPONENTS
c IN A CHEMICAL MIXTURE. IT IS ASSUMED THAT THE MIXTURE IS COMPRISED OF
c TWO COMPONENTS IN A FIXED MIXING RATIO. IT IS ALSO ASSUMED THAT A
c PROBIT MODEL ADEQUATELY DESCRIBES THE DOSE-RESPONSE RELATIONSHIP FOR
c EACH OF THE TWO CHEMICALS IF APPLIED ALONE. FOR A DESCRIPTION OF THE
c STATISTICAL ASPECTS OF THE PROCEDURE, SEE ROBERTSON AND SMITH,
c "JOINT ACTION OF PYRETHROIDS WITH ORGANOPHOSPHORUS AND CARBAMATE
c INSECTICIDES,...", J.ECON.ENT., FEBRUARY, 1984. THE NOTATION USED IN
c THE PROGRAM FOLLOWS THAT USED IN THE PAPER, WITH A FEW NECESSARY
c EXCEPTIONS.

c
c
c COMPANION PROGRAM IS PINORM.F77, A PROGRAM WHICH COMPUTES THE CDF OF
c A STANDARD NORMAL VARIABLE.

c
c

```
        implicit double precision(a-h,o-z)
        parameter(pi=3.141592654)
        character*10 file_in,file_out,chemical(2)
        dimension b1(2),b2(2),s1(2),s2(2),s12(2),adj(2),prop(2),
1  x(2),y(2),p(2),v(2),c(2)
```

c

```
        print *, 'enter file name where data is stored. '
        read *, file_in
        open(3,file=file_in)
        print *, 'enter file name to store output. '
        read *, file_out
        open(4,file=file_out)
```

c

```
1000    do 1 i=1,2
          read(3,*,end=9999) chemical(i)
          read(3,*) b1(i),b2(i),s1(i),s2(i),s12(i),adj(i),prop(i)
1      continue
```

c

```
        write(4,10) chemical,prop,b1,b2,s1,s2,s12,adj
10      format (1x,'chemical',a10,3x,a10/,
1          ' proportion',2f10.4,/,
2          ' intercept b1',2f10.4,/,
3          ' slope b2',2f10.4,/,
4          ' var(b1)',2f10.4,/,
5          ' var(b2)',2f10.4,/,
6          ' covar(b1,b2)',2f10.4,/,
7          ' het adjust',2f10.4)
```

c

```
        read(3,*) k, exp, dead
        control=dead/exp
        cc=control/((1.-control)*exp)
```

c

```
        write(4,20) chemical,exp,control
20      format(1x,'mixture of ',a10,' and ',a10,/,
1          ' dose no_exp obs.mort exp.mort chisq',/,
2          ' controls ',f6.0,f10.4)
        bigzsq=0.
        do 3 j=1,k
```

```

      read(3,*) dose,exp,dead
      phat=dead/exp
      do 2 i=1,2
         x(i)= prop(i)*dose
         y(i)=b1(i)+b2(i)*dlog10(x(i))
         p(i)=pinorm(y(i))
         v(i)=s1(i)+2.*s12(i)*dlog10(x(i))+s2(i)*dlog10(x(i))**2
         c(i)=dexp(-y(i)**2)*v(i)*adj(i)/(2.*pi*(1.-p(i))**2)
2      continue
      p0=control + (1.-control)*(p(1) + (1.-p(1))*p(2))
      var = phat*(1.-phat)/exp +
1      (1.-p0)**2*((1.+cc)*(1.+c(1))*(1.+c(2)) - 1.)
      zsq=(phat-p0)**2/var
30     write(4,30) dose,exp,phat,p0,zsq
      format(1x,f10.4,f6.0,3f10.4)
      bigzsq=bigzsq+zsq
3     continue
c
      write(4,40) bigzsq,k
40     format(1x,'chisq= ',f10.4,' with ',i4,' degrees of freedom'///)
c
      go to 1000
c
9999  stop
      end

```

DOUBLE PRECISION FUNCTION PINORM(Z)

```

C
C      THIS IS AN ADAPTATION ON THE SUBROUTINE PINORM PROGRAMMED
C      BY ROBERT RUSSELL AT FORT COLLINS.
C      THIS ROUTINE DOES NOT HAVE ENTRY POINTS FOR ERF AND ERFC
C
C      THE PROGRAM COMPUTES THE CUMULATIVE PROBABILITY FUNCTION
C      OF THE UNIT NORMAL DISTRIBUTION AT Z.
C
C      THE METHOD IS TO COMPUTE THE COMPLEMENTARY ERROR FUNCTION,
C      ERFC AND USE THE IDENTITY:  P(Z)=1 - .5*ERFC(Z/SQRT(2)).
C
C      FOR ABS(Z)/SQRT(2) GREATER THAN .47, USE HART'S RATIONAL
C      APPROXIMATION TO ERFC, THEN IDENTITY ABOVE.
C      FOR ABS(Z)/SQRT(2) LESS THAN .47, USE TAYLOR SERIES APPROX.
C      TO ERF, THEN ERFC=1-ERF, THEN IDENTITY ABOVE.
C
C      FINALLY, IF Z<0, P(Z)=1-P(-Z).
C
C      KS, DEC. 80

```

IMPLICIT DOUBLE PRECISION (A-Z)

```

DIMENSION A(4),B(5)
DATA A/10.0046412, 8.42655286, 3.46025933, .562353612/
DATA B/10.0046412, 19.7155807, 15.7022881, 6.09074879, 1./
X=DABS(Z)/DSQRT(2.D0)

```

```

      IF (X.GT..47) GO TO 20

```

```

      IF X<.47, USE TAYLOR SERIES APPROXIMATION.

```

```

W=X**2
E=X*(1.-W*((1./3.)-W*.5*((1./5.)-W*(1./3.)*((1./7.)-W*.25*
$ ((1./9.)-W*(1./55.)))))
E=1.-1.12837917*E
GO TO 40

```

C
C
C

```

FOR X>.47 BUT X<5. USE HART.

```

```

20 IF (X.GT.5.) GO TO 30
AA=A(1)+(A(2)+(A(3)+A(4)*X)*X)*X
BB=B(1)+(B(2)+(B(3)+(B(4)+B(5)*X)*X)*X)*X
E=EXP(-X**2)*AA/BB
GO TO 40

```

C
C
C

```

FOR X>5. RETURN ERFC=0.

```

```

30 E=0.

```

C
C
C

```

ADJUST FOR NEGATIVE ARGUMENTS.

```

```

40 IF (Z.LT.0.) E=2.-E
PINORM=1.-(.5*E)
RETURN
END

```

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